

HILBERT SPACE WITH REPRODUCING KERNEL AND UNIFORM DISTRIBUTION PRESERVING MAPS, II

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ABSTRACT. For Hilbert space H with reproducing kernel $K(\mathbf{x}, \mathbf{y})$, we express the mean square worst-case error

$$\int_{[0,1]^s} \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi(\{\mathbf{x}_n + \boldsymbol{\sigma}\})) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 d\boldsymbol{\sigma}$$

as

$$\frac{1}{N^2} \sum_{n,m=0}^{N-1} \int_{[0,1]^s} K(\Phi(\mathbf{x}), \Phi(\mathbf{y})) d\mathbf{x} d\mathbf{y} g_{m,n}(\mathbf{x}, \mathbf{y}) - \int_{[0,1]^{2s}} K(\mathbf{x}, \mathbf{y}) d\mathbf{x} d\mathbf{y},$$

where $\Phi(\mathbf{x})$ is a uniform distribution preserving map, $\mathbf{x}_0, \dots, \mathbf{x}_{N-1} \in [0, 1]^s$, and $g_{m,n}(\mathbf{x}, \mathbf{y})$ are copulas associated with points $\mathbf{x}_m$ and $\mathbf{x}_n$. Applying this, for dimension $s = 1$, we find that the minimum of the mean square worst-case error is attained in the sequence $x_n = \frac{n}{N}$, for the kernel $K(x, y) = 1 - \max(x, y)$ and $\Phi(x) = x$.

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1. Introduction

Let $\mathbf{x} \in [0, 1]^s$. Let H be a Hilbert space of some functions $f(\mathbf{x})$ from $L^2([0, 1]^s)$, with a reproducing kernel $K(\mathbf{x}, \mathbf{y})$ from $L^1([0, 1]^s)$, and let

$$\mathbf{x}_0, \dots, \mathbf{x}_{N-1} \in [0, 1]^s$$

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be an N -terms sequence. I.H. Sloan and H. Woźniakowski [11] proved that

$$\begin{aligned} & \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\mathbf{x}_n) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 \\ &= \int_{[0,1]^s} \int_{[0,1]^s} K(\mathbf{x}, \mathbf{y}) d\mathbf{x} d\mathbf{y} - \frac{2}{N} \sum_{n=0}^{N-1} \int_{[0,1]^s} K(\mathbf{x}_n, \mathbf{y}) d\mathbf{y} \\ &+ \frac{1}{N^2} \sum_{n,m=0}^{N-1} K(\mathbf{x}_m, \mathbf{x}_n), \end{aligned} \quad (1)$$

where the left hand side of (1) is said to be the worst-case square error. In the quasi-Monte Carlo (QMC) integration the following mean square worst-case error

$$\int_{[0,1]^s} \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\mathbf{x}_n \oplus \boldsymbol{\sigma}) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 d\boldsymbol{\sigma} \quad (2)$$

is studied for digital shift $\mathbf{x} \oplus \boldsymbol{\sigma}$. For a special map $\Phi(\mathbf{x})$ called the tent map and for some sequence $\mathbf{x}_n$ the following mean square worst-case error

$$\int_{[0,1]^s} \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi(\mathbf{x}_n \oplus \boldsymbol{\sigma})) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 d\boldsymbol{\sigma} \quad (3)$$

is better than (2), cf. [2]. In the previous paper [1] there is noted that the tent map $\Phi(\mathbf{x})$ and shift $\mathbf{x} \oplus \boldsymbol{\sigma}$ belong to the so called uniform distribution preserving maps. Here a map $\Phi : [0, 1]^s \rightarrow [0, 1]^s$ is called uniform distribution preserving (u.d.p.) map if for every uniformly distributed (u.d.) sequence $\mathbf{x}_n$, $n = 1, 2, \dots$, the image $\Phi(\mathbf{x}_n)$ is again a u.d. sequence. By Weyl limit [10, p. 1–62] we have the following main criterion of u.d.p. [10, 2.5.1].

THEOREM 1. *A map $\Phi(\mathbf{x})$ is u.d.p. if and only if $\Phi(\mathbf{x})$ is Riemann integrable and for every Riemann integrable $f : [0, 1]^s \rightarrow \mathbb{R}$ we have*

$$\int_{[0,1]^s} f(\Phi(\mathbf{x})) d\mathbf{x} = \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x}. \quad (4)$$

In [1] the sequence $\mathbf{x}_0, \dots, \mathbf{x}_{N-1}$ is replaced by $\Psi(\mathbf{x}_0, \boldsymbol{\sigma}), \dots, \Psi(\mathbf{x}_{N-1}, \boldsymbol{\sigma})$, where $\Psi(\mathbf{x}, \boldsymbol{\sigma})$ is a u.d.p. map with respect to $\mathbf{x}$ and $\boldsymbol{\sigma}$, simultaneously.

Applying (4) the mean square worst-case error (1) is expressed as

$$\begin{aligned} \int_{[0,1]^s} \sup_{\substack{f \in H \\ ||f|| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Psi(\mathbf{x}_n, \boldsymbol{\sigma})) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 d\boldsymbol{\sigma} = \\ \frac{1}{N^2} \sum_{n,m=0}^{N-1} \int_{[0,1]^s} K(\Psi(\mathbf{x}_m, \boldsymbol{\sigma}), \Psi(\mathbf{x}_n, \boldsymbol{\sigma})) d\boldsymbol{\sigma} - \int_{[0,1]^{2s}} K(\mathbf{x}, \mathbf{y}) d\mathbf{x} d\mathbf{y} \quad (5) \end{aligned}$$

because of (4), u.d.p. of $\Psi(\mathbf{x}, \boldsymbol{\sigma})$ implies

$$\int_{[0,1]^s} K(\Psi(\mathbf{x}_n, \boldsymbol{\sigma}), \mathbf{y}) d\boldsymbol{\sigma} = \int_{[0,1]^s} K(\boldsymbol{\sigma}, \mathbf{y}) d\boldsymbol{\sigma}$$

and thus

$$\frac{2}{N} \sum_{n=0}^{N-1} \int_{[0,1]^s} \int_{[0,1]^s} K(\Psi(\mathbf{x}_n, \boldsymbol{\sigma}), \mathbf{y}) d\mathbf{y} d\boldsymbol{\sigma} = 2 \int_{[0,1]^{2s}} K(\mathbf{x}, \mathbf{y}) d\mathbf{x} d\mathbf{y}. \quad (6)$$

By using Fourier-Walsh expansion of the kernel $K(\mathbf{x}, \mathbf{y})$

$$K(\mathbf{x}, \mathbf{y}) = \sum_{\mathbf{k}, \mathbf{k}' \in \mathbb{N}_0^s} \hat{K}(\mathbf{k}, \mathbf{k}') \overline{\text{wal}_{\mathbf{k}}(\mathbf{x})} \text{wal}_{\mathbf{k}'}(\mathbf{y}),$$

where

$$\hat{K}(\mathbf{k}, \mathbf{k}') = \int_{[0,1]^{2s}} K(\mathbf{x}, \mathbf{y}) \text{wal}_{\mathbf{k}}(\mathbf{x}) \overline{\text{wal}_{\mathbf{k}'}(\mathbf{y})} d\mathbf{x} d\mathbf{y} \quad (7)$$

in [1, Thm. 2] the following theorem is proved.¹

THEOREM 2. *For every sequence $\mathbf{x}_0, \dots, \mathbf{x}_{N-1}$ in the unit cube $[0,1]^s$ and every u.d.p. map $\Phi(\mathbf{x})$ and an arbitrary kernel $K(\mathbf{x}, \mathbf{y})$ with Fourier-Walsh expansion (7) we have*

$$\begin{aligned} \int_{[0,1]^s} \sup_{\substack{f \in H \\ ||f|| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi(\mathbf{x}_n \oplus \boldsymbol{\sigma})) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 d\boldsymbol{\sigma} = \\ \sum_{\substack{\mathbf{k} \in \mathbb{N}_0^s \\ \mathbf{k} \neq \mathbf{0}}} \hat{K}_1(\mathbf{k}, \mathbf{k}) \left| \frac{1}{N} \sum_{n=0}^{N-1} \text{wal}_{\mathbf{k}}(\mathbf{x}_n) \right|^2, \quad (8) \end{aligned}$$

where $\hat{K}_1(\mathbf{k}, \mathbf{k}) = \int_{[0,1]^{2s}} K(\Phi(\mathbf{x}), \Phi(\mathbf{y})) \overline{\text{wal}_{\mathbf{k}}(\mathbf{x})} \text{wal}_{\mathbf{k}}(\mathbf{y}) d\mathbf{x} d\mathbf{y}$.

¹Theorem 2 extend [2, Thm. 4] or [3, Thm. 12.7], which was proved for a Sobolev weighted space.

Theorem 2 separates the map $\Phi(\mathbf{x})$ and the sequence $\mathbf{x}_n$ in the mean square worst-case error (8).

The aim of this paper is to present a new expression of the mean square worst-case error by applying the theory of distribution functions (d.f.s) of sequences and the Riemann-Stieltjes integration. In Part 2, there is proved another type of separation as in (8), denoted by (9). Then in Part 3 we apply (9) in one-dimensional case for the special kernel $K(x, y) = 1 - \max(x, y)$ and for some extension $\Phi(x)$ of the tent map $\Phi(x) = 1 - |2x - 1|$. In Part 4 we also find the mean square worst-case error for $\Phi(x) = bx \bmod 1$. Finally, in Part 6, for the identity map $\Phi(x) = x$ we find that the minimum of the mean square worst-case error (9) over arbitrary points $x_0, x_1, \dots, x_{N-1}$ in $[0, 1]$ is attained at $x_i - x_{i-1} = \frac{1}{N}$.

2. Distribution functions method for dimension s

Following [1] we define:

- $\Psi(\mathbf{x}, \boldsymbol{\sigma})$ is a u.d.p. map of the form $\Phi(\mathbf{x} \oplus \boldsymbol{\sigma})$ or $\Phi(\{\mathbf{x} + \boldsymbol{\sigma}\})$, where $\Phi(\mathbf{x})$ is an arbitrary u.d.p. map.
- $x = \frac{x_0}{b} + \frac{x_1}{b^2} + \dots$ is a b -adic representation of $x \in [0, 1]$, and
- $\sigma = \frac{\sigma_0}{b} + \frac{\sigma_1}{b^2} + \dots$, then
- $x \oplus \sigma = \frac{x_0 + \sigma_0 \pmod{b}}{b} + \frac{x_1 + \sigma_1 \pmod{b^2}}{b^2} + \dots$,
- $\mathbf{x} \oplus \boldsymbol{\sigma} = (x_1 \oplus \sigma_1, x_2 \oplus \sigma_2, \dots, x_s \oplus \sigma_s)$.
- $\{\mathbf{x} + \boldsymbol{\sigma}\} = (\{x_1 + \sigma_1\}, \{x_2 + \sigma_2\}, \dots, \{x_s + \sigma_s\})$,
- $\boldsymbol{\sigma}_i, i = 0, 1, \dots$, is a u.d. sequence in $[0, 1]^s$,
- $g_{m,n}(\mathbf{x}, \mathbf{y})$ is the asymptotic distribution function (a.d.f.) of the sequence $(\mathbf{x}_m \oplus \boldsymbol{\sigma}_i, \mathbf{x}_n \oplus \boldsymbol{\sigma}_i), i = 0, 1, 2, \dots$,
- also the same notation $g_{m,n}(\mathbf{x}, \mathbf{y})$ is used for a.d.f. of the sequence $(\{\mathbf{x}_m + \boldsymbol{\sigma}_i\}, \{\mathbf{x}_n + \boldsymbol{\sigma}_i\}), i = 0, 1, 2, \dots$.
- We distinguish $g_{m,n}(\mathbf{x}, \mathbf{y})$ depending on whether $\Psi(\mathbf{x}, \boldsymbol{\sigma}) = \Phi(\mathbf{x} \oplus \boldsymbol{\sigma})$ or $\Psi(\mathbf{x}, \boldsymbol{\sigma}) = \Phi(\{\mathbf{x} + \boldsymbol{\sigma}\})$.
- Note that $g_{m,n}(\mathbf{x}, \mathbf{1}) = \mathbf{x}$ and $g_{m,n}(\mathbf{1}, \mathbf{y}) = \mathbf{y}$, and thus $g_{m,n}(\mathbf{x}, \mathbf{y})$ is a copula, see definition in [12, 2.3.].
- $\Phi(x) = 1 - |2x - 1|, x \in [0, 1]$, is u.d.p. map called the tent map or baker's map, see [5].
- Here, as usual, $\{x\}$ is the fractional part of x .

The following result holds for an arbitrary dimension s .

THEOREM 3. For every sequence $\mathbf{x}_0, \dots, \mathbf{x}_{N-1}$ in the unit cube $[0, 1]^s$ and every u.d.p. map $\Phi(\mathbf{x})$ and an arbitrary continuous kernel $K(\mathbf{x}, \mathbf{y})$ we have

$$\int_{[0,1]^s} \sup_{\substack{f \in H \\ ||f|| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Psi(\mathbf{x}_n, \boldsymbol{\sigma})) - \int_{[0,1]^s} f(\mathbf{x}) d\mathbf{x} \right|^2 d\boldsymbol{\sigma} = \frac{1}{N^2} \sum_{n,m=0}^{N-1} \int_{[0,1]^s} K(\Phi(\mathbf{x}), \Phi(\mathbf{y})) d\mathbf{x} d\mathbf{y} g_{m,n}(\mathbf{x}, \mathbf{y}) - \int_{[0,1]^{2s}} K(\mathbf{x}, \mathbf{y}) d\mathbf{x} d\mathbf{y}. \quad (9)$$

Proof. The integral (5) can be computed as limits

$$\begin{aligned} \frac{1}{M} \sum_{i=0}^{M-1} K(\Psi(\mathbf{x}_m, \boldsymbol{\sigma}_i), \Psi(\mathbf{x}_n, \boldsymbol{\sigma}_i)) &\rightarrow \int_{[0,1]^s} K(\Psi(\mathbf{x}_m, \boldsymbol{\sigma}), \Psi(\mathbf{x}_n, \boldsymbol{\sigma})) d\boldsymbol{\sigma}, \\ &\rightarrow \int_{[0,1]^s} K(\Phi(\mathbf{x}), \Phi(\mathbf{y})) d\mathbf{x} d\mathbf{y} g_{m,n}(\mathbf{x}, \mathbf{y}), \end{aligned}$$

where $M \rightarrow \infty$. Now, from (5), we have (9). $\square$

REMARK 4. Theorem 3 separates the map $\Phi(\mathbf{x})$ and the sequence $\mathbf{x}_n$ again by (9). Further an explicit formula of d.f. $g_{m,n}(\mathbf{x}, \mathbf{y})$ is given only for one-dimensional case in Part 5. Where it is used that

- $g_{m,n}(x, y) = |h_m^{-1}([0, x]) \cap h_n^{-1}([0, y])|$, where
- $h_n(\sigma) = x_n \oplus \sigma$ or $h_n(\sigma) = \{x_n + \sigma\}$ and

the Riemann-Stieltjes integral (9) is computed using integration by parts.

3. The case $s = 1$

THEOREM 5. Let $\Psi(x, \sigma)$ be $\Phi(x \oplus \sigma)$ or $\Phi(\{x + \sigma\})$, where $\Phi(x)$ is an arbitrary u.d.p. map. Let $x_0, \dots, x_{N-1}$ be a sequence in $[0, 1]$ and $K(x, y)$ be a continuous kernel. Then

$$\begin{aligned} \int_0^1 \sup_{\substack{f \in H \\ ||f|| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Psi(x_n, \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma = \\ \int_0^1 \int_0^1 \left(\frac{1}{N^2} \sum_{m,n=0}^{N-1} g_{m,n}(x, y) - xy \right) d_x d_y K(\Phi(x), \Phi(y)). \quad (10) \end{aligned}$$

Proof. To compute $\int_0^1 \int_0^1 K(\Phi(x), \Phi(y)) d_x d_y g_{m,n}(x, y)$, we use integration by parts

$$\begin{aligned} \int_0^1 \int_0^1 F(x, y) d_x d_y g(x, y) &= F(1, 1) - \int_0^1 g(1, y) d_y F(1, y) \\ &\quad - \int_0^1 g(x, 1) d_x F(x, 1) + \int_0^1 \int_0^1 g(x, y) d_x d_y F(x, y) \end{aligned} \quad (11)$$

which holds for an arbitrary continuous $F(x, y)$ and every d.f. $g(x, y)$.

REMARK 6. For every m, n we have

$$g_{m,n}(x, 1) = x \quad \text{and} \quad g_{m,n}(1, y) = y,$$

i.e. $g_{m,n}(x, y)$ is a two-dimensional copula. Every copula is continuous [7]. Furthermore by Fréchet-Hoeffding bounds

$$\max(x + y - 1, 0) \leq g_{m,n}(x, y) \leq \min(x, y) \quad \text{for} \quad (x, y) \in [0, 1]^2$$

and thus

$$\max(x + y - 1, 0) - xy \leq \frac{1}{N^2} \sum_{m,n=0}^{N-1} g_{m,n}(x, y) - xy \leq \min(x, y) - xy. \quad (12)$$

Proof of (11). Integration by parts gives

$$\begin{aligned} &\int_0^1 \int_0^1 F(x, y) d_x d_y g(x, y) \\ &= \left[\int_0^1 F(x, y) d_y g(x, y) \right]_{x=0}^{x=1} - \int_0^1 \int_0^1 d_y g(x, y) d_x F(x, y) \\ &= \int_0^1 F(1, y) d_y g(1, y) - \int_0^1 \int_0^1 d_y g(x, y) d_x F(x, y) \\ &= [F(1, y) g(1, y)]_{y=0}^{y=1} - \int_0^1 g(1, y) d_y F(1, y) \\ &\quad - \left[\int_0^1 g(x, y) d_x F(x, y) \right]_{y=0}^{y=1} \\ &\quad + \int_0^1 \int_0^1 g(x, y) d_y d_x F(x, y). \end{aligned}$$

Using $g(0, y) = g(x, 0) = 0$ for every $x, y \in [0, 1]$ we have (11). □

Now, applying (11) to (9) we find

$$\begin{aligned}
 & \int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Psi(x_n, \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma \\
 &= K(\Phi(1), \Phi(1)) - \int_0^1 y dy K(\Phi(1), \Phi(y)) \\
 &\quad - \int_0^1 x dx K(\Phi(x), \Phi(1)) \\
 &\quad + \int_0^1 \int_0^1 \left(\frac{1}{N^2} \sum_{m,n=0}^{N-1} g_{m,n}(x, y) \right) dx dy K(\Phi(x), \Phi(y)) \\
 &\quad - \int_0^1 \int_0^1 K(x, y) dx dy. \tag{13}
 \end{aligned}$$

From u.d.p. of $\Phi(x)$ and (11) follows that

$$\begin{aligned}
 \int_0^1 \int_0^1 K(x, y) dx dy &= \int_0^1 \int_0^1 K(\Phi(x), \Phi(y)) dx dy \\
 &= K(\Phi(1), \Phi(1)) - \int_0^1 y dy K(\Phi(1), \Phi(y)) \\
 &\quad - \int_0^1 x dx K(\Phi(x), \Phi(1)) \\
 &\quad + \int_0^1 \int_0^1 xy dx dy K(\Phi(x), \Phi(y)). \tag{14}
 \end{aligned}$$

Adding (13) and (14) we derive (10). $\square$

In the following theorem we apply Theorem 5, or precisely (13), to the new u.d.p. map called the tent map. The simple tent map is defined as

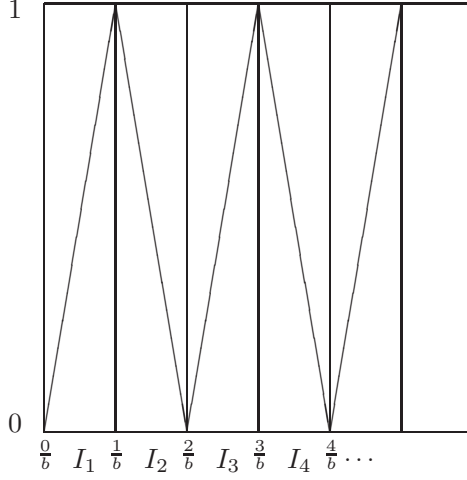
$$\Phi_0(x) = 1 - |2x - 1|.$$

We extend this map to the map $\Phi(x)$ with the graph in Fig. 1 and define by the following: putting

$$I_i = \left[\frac{i-1}{b}, \frac{i}{b} \right), \quad i = 1, 2, \dots, b,$$

we define

$$\Phi(x) = ((-1)^{i-1}(bx - i - 1)) + \frac{1}{2}(1 + (-1)^i) \text{ if } x \in I_i. \tag{15}$$


 FIGURE 1. The graph of the tent map $\Phi(x)$.

THEOREM 7. Let $K(x, y) = 1 - \max(x, y)$, $\Phi(x)$ be the tent u.d.p. map and $2|b$. Then we have

$$\begin{aligned}
 & \int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi(x_n \oplus \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma \\
 &= -\frac{1}{3} + \frac{1}{N^2} \sum_{m,n=0}^{N-1} \left[\sum_{\substack{i,j=1 \\ 2|i-j}}^b b \int_{\frac{i-1}{b}}^{\frac{i}{b}} g_{m,n} \left(x, x + \frac{j-i}{b} \right) dx \right. \\
 & \quad \left. - \sum_{\substack{i,j=1 \\ 2 \nmid i-j}}^b b \int_{\frac{i-1}{b}}^{\frac{i}{b}} g_{m,n} \left(x, -x + \frac{i+j-1}{b} \right) dx \right]. \tag{16}
 \end{aligned}$$

The same holds also for $\Phi(\{x_n + \sigma\})$ instead of $\Phi(x_n \oplus \sigma)$. The base b of the tent u.d.p. map $\Phi(x)$ is independent of the base of the shift $x_n \oplus \sigma$.

Proof. The Riemann-Stieltjes integral $\int_0^1 \int_0^1 g(x, y) d_x d_y F(x, y)$ is defined as the limit

$$\begin{aligned}
 & \sum_{k=1}^m \sum_{l=1}^n g(\alpha_k, \beta_l) (F(x_k, y_l) + F(x_{k+1}, y_{l+1}) - F(x_k, y_{l+1}) - F(x_{k+1}, y_l)) \rightarrow \\
 & \int_0^1 \int_0^1 g(x, y) d_x d_y F(x, y) \tag{17}
 \end{aligned}$$

if the maximal diameter of the rectangles $[x_k, x_{k+1}] \times [y_l, y_{l+1}]$ tends to zero. This integral exists for a continuous $g(x, y)$ and for $F(x, y)$ with a bounded variation. In the following we assume that these conditions are valid. For partitions $0 = x_0 < x_1 < \dots < x_m = 1$ of x -axis and $0 = y_0 < y_1 < \dots < y_n = 1$ of y -axis we assume that they are the same and contain $\frac{i}{b}$, $i = 1, 2, \dots, b$. Denote the differential $d_x d_y F(x, y)$ for the rectangle $\square = [x_k, x_{k+1}] \times [y_l, y_{l+1}]$ by

$$\square F(x, y) = F(x_k, y_l) + F(x_{k+1}, y_{l+1}) - F(x_k, y_{l+1}) - F(x_{k+1}, y_l). \quad (18)$$

If the diameter of $\square$ tends to zero, $\square \rightarrow 0$, then we find the differential $d_x d_y F(x, y)$ as $d_x d_y F(x, y) = F(x, y) + F(x + dx, y + dy) - F(x, y + dy) - F(x + dx, y)$.

By the definition of Riemann-Stieltjes integral

$$\int_0^1 \int_0^1 g(x, y) d_x d_y F(x, y) = \sum_{i,j=1}^b \iint_{I_i \times I_j} g(x, y) d_x d_y F(x, y).$$

In the following we put $F(x, y) = \max(\Phi(x), \Phi(y))$ and distinguish four cases:
 1⁰. i, j -are even, $x \in I_i$ and $y \in I_j$, see Fig. 2.

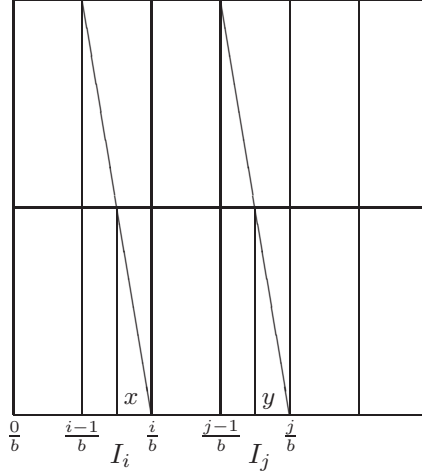


FIGURE 2. The graph of the tent map $\Phi(x)$ in the case 1⁰.

Then we have

$$\Phi(x) = \Phi(y) \Leftrightarrow b \left(\frac{i}{b} - x \right) = b \left(\frac{j}{b} - y \right) \Leftrightarrow y = x + \frac{j-i}{b}$$

and

$$\Phi(x) = \Phi(y) > \Phi(y + dy).$$

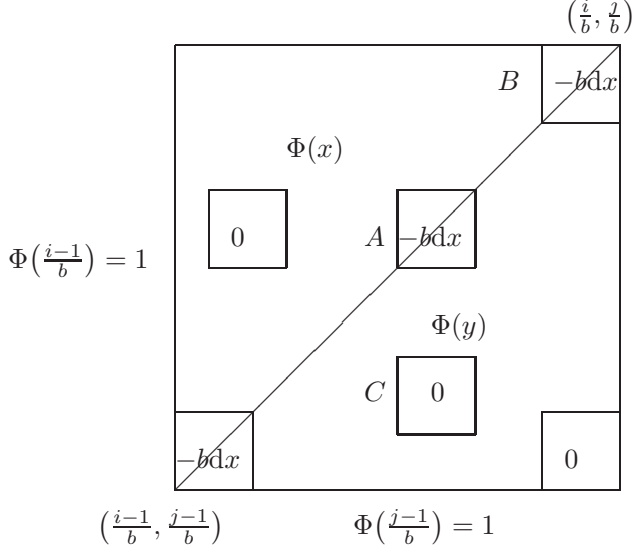


FIGURE 3.

In Fig. 3 we have the differential $F(x, y)$ in $I_i \times I_j$ in the case 1^0 , where the differential $\square F(x, y)$ for squares $\square = A, B, C$, respectively, is computed by means of definition (18):

$$\begin{aligned}
 AF(x, y) &= (\Phi(x_k) = \Phi(y_l)) + (\Phi(x_{k+1}) = \Phi(y_{l+1})) - \Phi(x_k) - \Phi(y_l) \\
 &= \Phi(x_{k+1}) - \Phi(x_k) = -b(x_{k+1} - x_k), \\
 BF(x, y) &= (\Phi(x_k) = \Phi(y_l)) + (\Phi(x_{k+1}) = \Phi(y_{l+1}) = 0) - \Phi(x_k) - \Phi(y_l) \\
 &= \Phi(x_{k+1}) - \Phi(x_k) = -b(x_{k+1} - x_k), \\
 CF(x, y) &= \Phi(y_l) + \Phi(y_{l+1}) - \Phi(y_l) - \Phi(y_{l+1}) = 0.
 \end{aligned}$$

From this

$$\iint_{I_i \times I_j} g(x, y) \mathrm{d}_x \mathrm{d}_y F(x, y) = \int_{\frac{i-1}{b}}^{\frac{i}{b}} g\left(x, x + \frac{j-i}{b}\right) (-b) \mathrm{d}x. \quad (19)$$

Similarly,

2^0 . i, j -are odd, $x \in I_i$ and $y \in I_j$, see Fig. 4.

HILBERT SPACE WITH REPRODUCING KERNEL

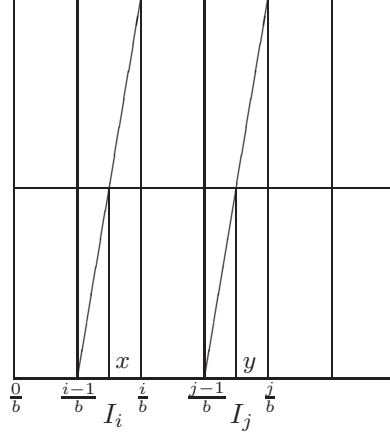


FIGURE 4. The graph of $\Phi(x)$ in the case 2^0 .

$$\Phi(x) = \Phi(y) \Leftrightarrow b\left(x - \frac{i-1}{b}\right) = b\left(y - \frac{j-1}{b}\right) \Leftrightarrow y = x + \frac{j-i}{b}$$

and

$$\Phi(x) = \Phi(y) < \Phi(y + dy).$$

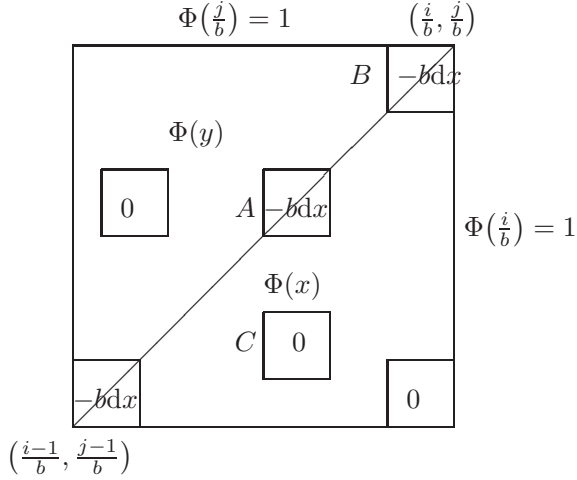


FIGURE 5.

In Fig. 5 we have the differential $F(x, y)$ in $I_i \times I_j$ in the case 2^0 , where the differential $\square F(x, y)$ for squares $\square = A, B, C$, respectively, is computed

by means of definition (18):

$$\begin{aligned} AF(x, y) &= (\Phi(x_k) = \Phi(y_l)) + (\Phi(x_{k+1}) = \Phi(y_{l+1})) - \Phi(y_{l+1}) - \Phi(x_{k+1}) \\ &= \Phi(x_k) - \Phi(x_{k+1}) = -b(x_{k+1} - x_k), \end{aligned}$$

$$BF(x, y) = (\Phi(x_k) = \Phi(y_l)) + 1 - 1 - 1 = \Phi(x_k) - \Phi(1) = -b(1 - x_k),$$

$$CF(x, y) = \Phi(x_k) + \Phi(x_{k+1}) - \Phi(x_k) - \Phi(x_{k+1}) = 0.$$

Similarly, for other $I_i \times I_j$. Then for Riemann-Stieltjes integral

$$\iint_{I_i \times I_j} g(x, y) d_x d_y F(x, y) = \int_{\frac{i-1}{b}}^{\frac{i}{b}} g\left(x, x + \frac{j-i}{b}\right) (-b) dx. \quad (20)$$

3⁰. i -even and j -odd, and $x \in I_i$ and $y \in I_j$, see Fig. 6.

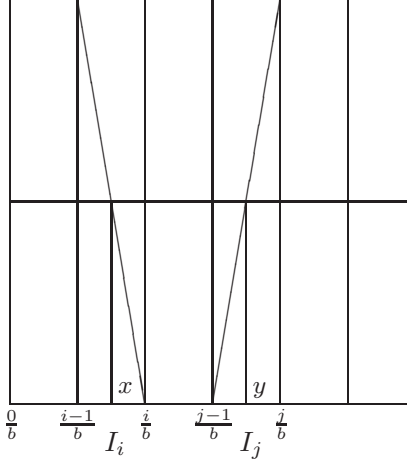


FIGURE 6. The graph of $\Phi(x)$ in the case 3⁰.

$$\begin{aligned} \Phi(x) = \Phi(y) &\Leftrightarrow b\left(\frac{i}{b} - x\right) = b\left(y - \frac{j-1}{b}\right) \Leftrightarrow y = -x + \frac{i+j-1}{b} \\ \text{and} \\ \Phi(x) = \Phi(y) &< \Phi(y + dy). \end{aligned}$$

HILBERT SPACE WITH REPRODUCING KERNEL

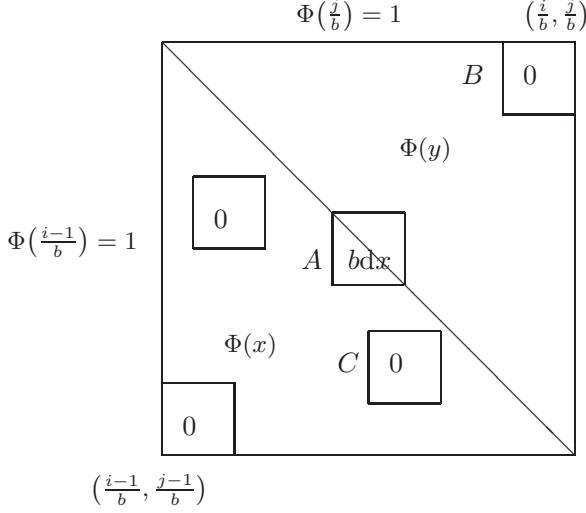


FIGURE 7.

In Fig. 7 we have the differential $F(x, y)$ in $I_i \times I_j$ in the case 3^0 , where the differential $\square F(x, y)$ for squares $\square = A, B, C$, respectively, is computed by means of definition (18):

$$\begin{aligned} AF(x, y) &= \Phi(x_k) + \Phi(y_{l+1}) - (\Phi(x_k) = \Phi(y_{l+1})) - (\Phi(x_{k+1}) = \Phi(y_l)) \\ &= \Phi(x_k) - \Phi(x_{k+1}) = b(x_{k+1} - x_k), \end{aligned}$$

$$BF(x, y) = \Phi(y_l) + 1 - 1 - \Phi(y_l) = 0,$$

$$CF(x, y) = \Phi(x_k) + \Phi(x_{k+1}) - \Phi(x_k) - \Phi(x_{k+1}) = 0.$$

From this

$$\iint_{I_i \times I_j} g(x, y) dx dy F(x, y) = \int_{\frac{i-1}{b}}^{\frac{i}{b}} g\left(x, -x + \frac{i+j-1}{b}\right) b dx. \quad (21)$$

Similarly,

4^0 . i -odd and j -even.

$$\iint_{I_i \times I_j} g(x, y) dx dy F(x, y) = \int_{\frac{i-1}{b}}^{\frac{i}{b}} g\left(x, -x + \frac{i+j-1}{b}\right) b dx. \quad (22)$$

For an application of (13) we note that

$$\Phi(1) = 0, \quad K(\Phi(1), \Phi(1)) = 1, \quad K(\Phi(x), \Phi(1)) = 1 - \Phi(x),$$

$$\begin{aligned} - \int_0^1 x d_x K(\Phi(x), \Phi(1)) &= \int_0^1 x d_x \Phi(x) \\ &= \int_0^{\frac{1}{b}} x b dx + \int_{\frac{1}{b}}^{\frac{2}{b}} x(-b) dx + \int_{\frac{2}{b}}^{\frac{3}{b}} x b dx + \cdots = -\frac{1}{2}, \end{aligned}$$

$$\int_0^1 \int_0^1 K(x, y) dx dy = \frac{1}{3}, \quad g(x, y) = \frac{1}{N^2} \sum_{m, n=0}^{N-1} g_{m, n}(x, y),$$

$$K(\Phi(x), \Phi(y)) = 1 - F(x, y).$$

Then we find (16). □

4. Further generalization of tent map

To study u.d.p. map $\Phi(x) = bx \bmod 1$ we use the further generalization of tent u.d.p. map defined in (15). We extended it to a two-parametric u.d.p. tent map $\Phi_{b_1, b_2}(x)$ defined by the graph in Fig. 8 that is also u.d.p. map.

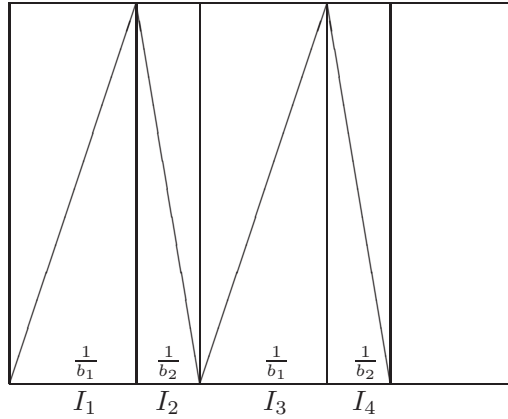


FIGURE 8. $\Phi_{b_1, b_2}(x)$ -tent function.

Here note that

$$I_i = \begin{cases} \left[\frac{i}{k} - \frac{1}{b_2}, \frac{i}{k} \right] & \text{if } i \text{ is even,} \\ \left[\frac{i-1}{k}, \frac{i-1}{k} + \frac{1}{b_1} \right] & \text{if } i \text{ is odd.} \end{cases}$$

Then, similarly to (16), the mean square worst-case error can be expressed as

$$\begin{aligned} & \int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi_{b_1, b_2}(x_n \oplus \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma \\ &= -\frac{1}{3} + \frac{1}{N^2} \sum_{m, n=0}^{N-1} \left[\sum_{\substack{i, j=1 \\ i, j \text{ even}}}^k b_2 \int_{\frac{i}{k} - \frac{1}{b_2}}^{\frac{i}{k}} g_{m, n} \left(x, x + \frac{j-i}{k} \right) dx \right. \\ & \quad + \sum_{\substack{i, j=1 \\ i, j \text{ odd}}}^k b_1 \int_{\frac{i-1}{k}}^{\frac{i-1}{k} + \frac{1}{b_1}} g_{m, n} \left(x, x + \frac{j-i}{k} \right) dx \\ & \quad - \sum_{\substack{i, j=1 \\ i \text{ even}, j \text{ odd}}}^k b_2 \int_{\frac{i}{k} - \frac{1}{b_2}}^{\frac{i}{k}} g_{m, n} \left(x, \frac{b_2}{b_1} \left(\frac{i}{k} - x \right) + \frac{j-1}{k} \right) dx \\ & \quad \left. - \sum_{\substack{i, j=1 \\ i \text{ odd}, j \text{ even}}}^k b_1 \int_{\frac{i-1}{k}}^{\frac{i-1}{k} + \frac{1}{b_1}} g_{m, n} \left(x, \frac{b_1}{b_2} \left(\frac{i-1}{k} - x \right) + \frac{j}{k} \right) dx \right], \quad (23) \end{aligned}$$

where $K(x, y) = 1 - \max(x, y)$, $\Phi_{b_1, b_2}(x)$ is a two-parameters tent u.d.p. map, $\frac{2}{k} = \frac{1}{b_1} + \frac{1}{b_2}$ and $2|k$. This also holds for $\Phi(\{x_n + \sigma\})$ instead of $\Phi(x_n \oplus \sigma)$. The bases b_1, b_2 for the u.d.p. map $\Phi_{b_1, b_2}(x)$ are independent of the base of the shift $x_n \oplus \sigma$. The proof of (23) is similar to the proof of Theorem 7.

Let $\Phi(x) = bx \bmod 1$ be the u.d.p. map and $k = 2b$. Since $K(\Phi(x), \Phi(y))$ for $\Phi(x) = bx \bmod 1$ is discontinuous, we cannot use (11). But we can use limits

$$\lim_{b_2 \rightarrow \infty} \Phi_{b_1, b_2}(x_m \oplus \sigma) = \Phi(x_m \oplus \sigma) \text{ on } [0, 1],$$

or

$$\lim_{b_2 \rightarrow \infty} \Phi_{b_1, b_2}(\{x_m + \sigma\}) = \Phi(\{x_m + \sigma\}),$$

as $b_2 \rightarrow \infty$, $k = \text{constant}$, $b_1 \rightarrow b = \frac{k}{2}$. From (23), the following limits hold:

$$\begin{aligned} & b_2 \int_{\frac{i}{k} - \frac{1}{b_2}}^{\frac{i}{k}} g_{m, n} \left(x, x + \frac{j-i}{k} \right) dx \rightarrow g_{m, n} \left(\frac{i}{k}, \frac{j}{k} \right), \\ & b_1 \int_{\frac{i-1}{k}}^{\frac{i-1}{k} + \frac{1}{b_1}} g_{m, n} \left(x, x + \frac{j-i}{k} \right) dx \rightarrow \frac{k}{2} \int_{\frac{i-1}{k}}^{\frac{i}{k} + \frac{2}{k}} g_{m, n} \left(x, x + \frac{j-i}{k} \right) dx, \end{aligned}$$

$$\begin{aligned}
& b_2 \int_{\frac{i}{k} - \frac{1}{b_2}}^{\frac{i}{k}} g_{m,n} \left(x, \frac{b_2}{b_1} \left(\frac{i}{k} - x \right) + \frac{j-1}{k} \right) dx = \\
& b_1 \int_{\frac{j-1}{k}}^{\frac{j-1}{k} + \frac{1}{b_1}} g_{m,n} \left(\frac{i}{k} + \frac{b_1}{b_2} \left(\frac{j-1}{k} - y \right), y \right) dy \rightarrow \frac{k}{2} \int_{\frac{j-1}{k}}^{\frac{j-1}{k} + \frac{2}{k}} g_{m,n} \left(\frac{i}{k}, x \right) dx, \\
& b_1 \int_{\frac{j-1}{k}}^{\frac{j-1}{k} + \frac{1}{b_1}} g_{m,n} \left(x, \frac{b_1}{b_2} \left(\frac{i-1}{k} - x \right) + \frac{j}{k} \right) dx \rightarrow \frac{k}{2} \int_{\frac{j-1}{k}}^{\frac{j-1}{k} + \frac{2}{k}} g_{m,n} \left(x, \frac{j}{k} \right) dx.
\end{aligned}$$

Finally, the limit

$$\begin{aligned}
& \int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi_{b_1, b_2}(x_n \oplus \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma \rightarrow \\
& \int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi(x_n \oplus \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma
\end{aligned}$$

follows from the expression (5) and from the limit

$$\int_0^1 K(\Phi_{b_1, b_2}(x_m \oplus \sigma), \Phi_{b_1, b_2}(x_n \oplus \sigma)) d\sigma \rightarrow \int_0^1 K(\Phi(x_m \oplus \sigma), \Phi(x_n \oplus \sigma)) d\sigma$$

which follows from the Lebesgue theorem on the dominated convergence. Then we have:

THEOREM 8. *Let $K(x, y) = 1 - \max(x, y)$ be the kernel, let $\Phi(x) = bx \bmod 1$ be the u.d.p. map and $k = 2b$. Then we have*

$$\begin{aligned}
& \int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\Phi(x_n \oplus \sigma)) - \int_0^1 f(x) dx \right|^2 d\sigma \\
& = -\frac{1}{3} + \frac{1}{N^2} \sum_{m,n=0}^{N-1} \left[\sum_{\substack{i,j=1 \\ i,j-\text{even}}}^k g_{m,n} \left(\frac{i}{k}, \frac{j}{k} \right) + \sum_{\substack{i,j=1 \\ i,j-\text{odd}}}^k \frac{k}{2} \int_{\frac{i-1}{k}}^{\frac{i}{k} + \frac{2}{k}} g_{m,n} \left(x, x + \frac{j-i}{k} \right) dx \right. \\
& \quad \left. - \sum_{\substack{i,j=1 \\ i-\text{even}, j-\text{odd}}}^k \frac{k}{2} \int_{\frac{j-1}{k}}^{\frac{j-1}{k} + \frac{2}{k}} g_{m,n} \left(\frac{i}{k}, x \right) dx - \sum_{\substack{i,j=1 \\ i-\text{odd}, j-\text{even}}}^k \frac{k}{2} \int_{\frac{i-1}{k}}^{\frac{i-1}{k} + \frac{2}{k}} g_{m,n} \left(x, \frac{j}{k} \right) dx \right].
\end{aligned} \tag{24}$$

The same holds also for $\Phi(\{x_n + \sigma\})$ instead of $\Phi(x_n \oplus \sigma)$. The base b for this u.d.p. map $\Phi(x)$ is independent of the base of the shift $x_n \oplus \sigma$.

5. Explicit formula for copulas $g_{m,n}(x, y)$

To apply Theorems 5 and 7 we need an expression of $g_{m,n}(x, y)$. For the sequence $(\{x_m + \sigma_i\}, \{x_n + \sigma_i\})$, $i = 1, 2, \dots$. Let $0 \leq u \leq v \leq 1$ be fixed and $x, y \in [0, 1]$ be variables and define:

$$h_u(x) = \{x + u\}, \quad h_v(y) = \{y + v\},$$

$$H(u, v, x, y) = |h_u^{-1}([0, x]) \cap h_v^{-1}([0, y])|.$$

Then we have

$$H(x_m, x_n, x, y) = g_{m,n}(x, y).$$

Using graphs of $h_u(x)$ and $h_v(y)$,

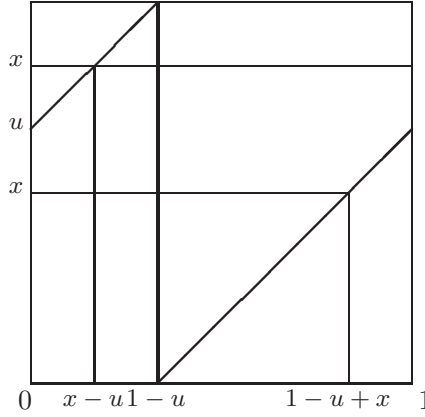


FIGURE 9. The graph of $h_u(x)$.

then we see

$$h_u^{-1}([0, x]) = \begin{cases} [1-u, 1-u+x] & \text{if } x \leq u, \\ [0, x-u] \cup [1-u, 1] & \text{if } u \leq x. \end{cases} \quad (25)$$

Hence

$$H(u, v, x, y) = \begin{cases} |[1-u, 1-u+x] \cap [1-v, 1-v+y]| & \text{if } x \leq u, y \leq v, \\ |[1-u, 1-u+x] \cap ([0, y-v] \cup [1-v, 1])| & \text{if } x \leq u, y > v, \\ |[0, x-u] \cup [1-u, 1] \cap [1-v, 1-v+y]| & \text{if } x > u, y \leq v, \\ |[0, x-u] \cup [1-u, 1] \cap ([0, y-v] \cup [1-v, 1])| & \text{if } x > u, y > v. \end{cases} \quad (26)$$

Now we are using minimum and maximum formulae for the length of intersection of two intervals $[\alpha, \beta]$ and $[\gamma, \delta]$

$$|[\alpha, \beta] \cap [\gamma, \delta]| = \max(\min(\beta, \delta) - \max(\alpha, \gamma), 0). \quad (27)$$

Insert (27) into (26), and we see that

$$H(u, v, x, y) = \begin{cases} \max(\min(y, x - u + v), 0) & \text{if } x \leq u, y \leq v, \\ \max(\min(x, y - v - 1 + u), 0) + \max(v - u + x, 0) & \text{if } x \leq u, y > v, \\ y & \text{if } x > u, y \leq v, \\ \min(x - u + v, y) + \max(y - v - 1 + u, 0) & \text{if } x > u, y > v. \end{cases} \quad (28)$$

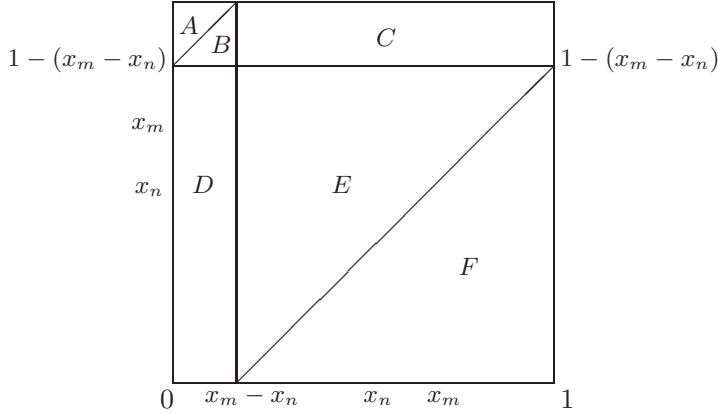


FIGURE 10. Division of $[0, 1]^2$.

Using division in Fig. 10 we have

$$H(u, v, x, y) = \begin{cases} x & \text{if } (x, y) \in A, \\ y - (1 - (u - v)) & \text{if } (x, y) \in B, \\ x + y - 1 & \text{if } (x, y) \in C, \\ 0 & \text{if } (x, y) \in D, \\ x - (u - v) & \text{if } (x, y) \in E, \\ y & \text{if } (x, y) \in F. \end{cases} \quad (29)$$

Since for a u.d. sequence $u_n \in [0, 1)$, $n = 1, 2, \dots$, the two-dimensional sequence (u_n, u_n) has an a.d.f. $g(x, y) = \min(x, y)$, then $H(u, v, x, y) = \min(x, y)$ for $u = v$. It can also be seen from (29).

Taken together we get $g_{m,m}(x, y) = \min(x, y)$ and for $x_m \neq x_n$ we have

$$g_{m,n}(x, y) = \begin{cases} x & \text{if } (x, y) \in A, \\ y - (1 - |x_m - x_n|) & \text{if } (x, y) \in B, \\ x + y - 1 & \text{if } (x, y) \in C, \\ 0 & \text{if } (x, y) \in D, \\ x - |x_m - x_n| & \text{if } (x, y) \in E, \\ y & \text{if } (x, y) \in F. \end{cases} \quad (30)$$

6. Application of Theorem 3

In this part we find global minimum of the mean square worst-case error for dimension $s = 1$, kernel $K(x, y) = 1 - \max(x, y)$ and $\Phi(x) = x$.

THEOREM 9. *Let $K(x, y) = 1 - \max(x, y)$, $\Phi(\{x + \sigma\}) = \{x + \sigma\}$, $0 \leq x_0 < x_1 < \dots < x_{N-2} < x_{N-1} \leq 1$ and put $t_i = x_{i+1} - x_i$, $i = 0, 1, \dots, N - 2$. Then the minimum of mean square worst-case error (9) is attained at $t_i = \frac{1}{N}$, $i = 0, 1, \dots, N - 2$ and this minimum is*

$$\int_0^1 \sup_{\substack{f \in H \\ \|f\| \leq 1}} \left| \frac{1}{N} \sum_{n=0}^{N-1} f(\{x_n + \sigma\}) - \int_0^1 f(x) dx \right|^2 d\sigma = \frac{1}{6N^2}. \quad (31)$$

Proof. We start with the formula (9). The following proof is divided into the parts 1–5 according to the following.

In the part 1 we shall express the integral (9) for the case $s = 1$ as

$$\begin{aligned} & \int_0^1 \int_0^1 K(\Phi(x), \Phi(y)) dx dy g_{m,n}(x, y) \\ &= \int_0^{|x_m - x_n|} K(\Phi(x), \Phi(x + 1 - |x_m - x_n|)) dx \\ &+ \int_{|x_m - x_n|}^1 K(\Phi(x), \Phi(x - |x_m - x_n|)) dx. \end{aligned} \quad (32)$$

In the part 2, for $\Phi(x) = x$, and $T_{m,n} = |x_m - x_n| = \sum_{k=m}^{n-1} t_k$, $m < n$, we shall prove

$$\frac{\partial}{\partial t_i} \left(\int_0^{T_{m,n}} K(x, x+1-T_{m,n}) dx + \int_{T_{m,n}}^1 K(x, x-T_{m,n}) dx \right) = 2T_{m,n} - 1 \quad (33)$$

assuming that $T_{m,n}$ contains term t_i .

In the part 3 we shall see that the zero partial derivative

$$\frac{\partial}{\partial t_i} \left(\sum_{\substack{m < n \\ m, n=0}}^{N-1} \int_0^1 \int_0^1 K(x, y) d_x d_y g_{m,n}(x, y) \right) = \sum_{\substack{m < n, m, n=0 \\ T_{m,n} \text{ contains } t_i}}^{N-1} (2T_{m,n} - 1) = 0 \quad (34)$$

is equivalent to the following linear equation

$$\begin{aligned} & t_0(N-2-i+1) + t_1 2(N-2-i+1) + \dots \\ & \dots + t_k(k+1)(N-2-i+1) + \dots \\ & \dots + t_i(i+1)(N-2-i+1) + t_{i+1}(i+1)(N-2-(i+1)+1) + \dots \\ & \dots + t_s(i+1)(N-2-s+1) + \dots \\ & \dots + t_{N-2}(i+1) = \frac{1}{2}(i+1)(N-2-i+1). \end{aligned} \quad (35)$$

In the part 4 we shall prove that the system (35), $i = 0, 1, \dots, N-2$, is regular and has the unique solution $t_i = \frac{1}{N}$ for $i = 0, 1, \dots, N-2$.

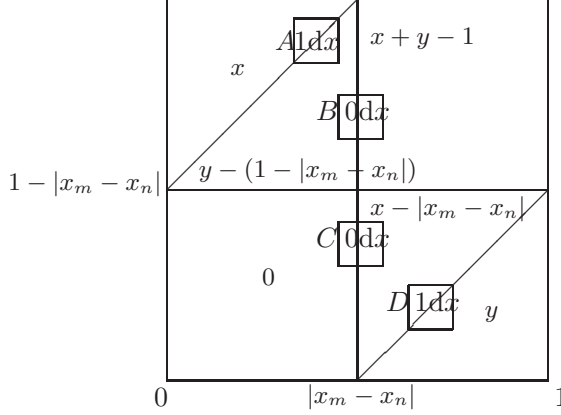
In the final part 5 we shall compute the minimum of the mean square worst-case error for $t_i = \frac{1}{N}$.

Proof of 1. In this case we calculate (9) directly by Riemann-Stieltjes integration. As in (18) we denote the differential $d_x d_y g_{m,n}(x, y)$ for the rectangle $\square = [x_k, x_{k+1}] \times [y_l, y_{l+1}]$ as

$$\begin{aligned} \square g_{m,n}(x, y) &= g_{m,n}(x_k, y_l) + g_{m,n}(x_{k+1}, y_{l+1}) \\ &\quad - g_{m,n}(x_k, y_{l+1}) - g_{m,n}(x_{k+1}, y_l). \end{aligned} \quad (36)$$

If diameter $\square \rightarrow 0$, then we find the differential $d_x d_y g_{m,n}(x, y)$ as

$$\begin{aligned} d_x d_y g_{m,n}(x, y) &= g_{m,n}(x, y) + g_{m,n}(x+dx, y+dy) \\ &\quad - g_{m,n}(x, y+dy) - g_{m,n}(x+dx, y). \end{aligned}$$


 FIGURE 11. Differential of $g_{m,n}(x, y)$.

From Fig. 11 we obtain:

$$\begin{aligned}
 Ag_{m,n}(x, y) &= (x_k = y_l - (1 - |x_m - x_n|)) + (x_{k+1} = y_{l+1} - (1 - |x_m - x_n|)) \\
 &\quad - x_k - (y_l - (1 - |x_m - x_n|)) \\
 &= y_l - (1 - |x_m - x_n|) + x_{k+1} - x_k - (y_l - (1 - |x_m - x_n|)) \\
 &= (x_{k+1} - x_k) \rightarrow 1 \cdot dx,
 \end{aligned}$$

$$\begin{aligned}
 Bg_{m,n}(x, y) &= (y_l - (1 - |x_m - x_n|)) + (x_{k+1} + y_{l+1} - 1) \\
 &\quad - (y_{l+1} - (1 - |x_m - x_n|)) - (x_{k+1} + y_l - 1) = 0,
 \end{aligned}$$

$$Cg_{m,n}(x, y) = 0 + (x_{k+1} - |x_m - x_n|) - 0 - (x_{k+1} - |x_m - x_n|) = 0,$$

$$\begin{aligned}
 Dg_{m,n}(x, y) &= (x_k - |x_m - x_n| = y_l) + (x_{k+1} - |x_m - x_n| = y_{l+1}) \\
 &\quad - (x_k - |x_m - x_n|) - y_l \\
 &= y_l + (x_{k+1} - |x_m - x_n|) - (x_k - |x_m - x_n|) - y_l \\
 &= (x_{k+1} - x_k) \rightarrow 1 \cdot dx.
 \end{aligned}$$

Thus the differential $d_x d_y g_{m,n}(x, y)$ is nonzero only on two lines:

- (i) $y = x + (1 - |x_m - x_n|)$ if $x \in [0, |x_m - x_n|]$, and
- (ii) $y = x - |x_m - x_n|$ if $x \in [|x_m - x_n|, 1]$.

Then (i) and (ii) imply (32). This holds for an arbitrary kernel satisfying (1) and an arbitrary u.d.p. function $\Phi(x)$. $\square$

Proof of 2. Here we assume that $K(x, y) = 1 - \max(x, y)$ and $\Phi(x) = x$. We have, for $t + T \in [0, 1]$, t is a variable and T is a constant

$$\begin{aligned} \frac{d}{dt} \int_0^{t+T} K(x, x+1-t) dx &= \frac{d}{dt} \int_{t+T}^{t+T+dt} K(x, x+1-(t+T)) dx + \\ &\quad \int_0^{t+T+dt} \frac{K(x, x+1-(t+T+dt)) - K(x, x+1-(t+T))}{dt} dx, \\ K(t+T, 1) + \int_0^{t+T} \left[\frac{\partial K(x, y)}{\partial y} \right]_{y=x+1-(t+T)} \frac{d}{dt} (x+1-(t+T)) dx &= \\ &= 0 + t + T. \end{aligned} \quad (37)$$

Similarly,

$$\begin{aligned} \frac{d}{dt} \int_{t+T}^1 K(x, x-(t+T)) dx \\ &= -K(t+T, 0) + \int_{t+T}^1 \left[\frac{\partial K(x, y)}{\partial y} \right]_{y=x-(t+T)} \frac{d}{dt} (x-(t+T)) dx \\ &= -1 + t + T + 0. \end{aligned} \quad (38)$$

Now, (37) and (38) imply (33). $\square$

The (34) implies (35) bearing in mind the following number of

$$\begin{aligned} T_{m,n}, 0 \leq m < n \leq N-1; \\ \#\{T_{m,n} \text{ contains } t_0, t_i\} &= 1(N-2-i+1), \\ \#\{T_{m,n} \text{ contains } t_1, t_i\} &= 2(N-2-i+1), \\ \#\{T_{m,n} \text{ contains } t_k, t_i\} &= (k+1)(N-2-i+1), \quad k < i, \\ \#\{T_{m,n} \text{ contains } t_i\} &= (i+1)(N-2-i+1), \\ \#\{T_{m,n} \text{ contains } t_{i+1}, t_i\} &= (i+1)(N-2-(i+1)+1), \\ \#\{T_{m,n} \text{ contains } t_s, t_i\} &= (i+1)(N-2-s+1), \quad s > i, \\ \#\{T_{m,n} \text{ contains } t_{N-2}, t_i\} &= (i+1), \\ \#\{T_{m,n}\} &= (N-2+1) + (N-2) + \cdots + 1 = \frac{(N-1)N}{2}. \end{aligned}$$

$\square$

Proof of 4. Put in (35)

$$t_0 = t_1 = \cdots = t_k = \cdots = t_i = \cdots = t_s = \cdots = t_{N-2} = 1.$$

Then we have

$$\begin{aligned} & (N-2-i+1) + 2(N-2-i+1) + \cdots + (k+1)(N-2-i+1) + \cdots \\ & + (i+1)(N-2-i+1) \\ & + (i+1)(N-2-(i+1)+1) + \cdots + (i+1)(N-2-s+1) + \cdots + (i+1) \\ & = \frac{(i+1)(i+2)}{2}(N-2-i+1) + (i+1)\frac{(N-2-i)(N-2-i+1)}{2} \\ & = \frac{(i+1)}{2}(N-2-i+1)(i+2+N-2-i). \end{aligned}$$

Thus

$$t_0 = t_1 = \cdots = t_k = \cdots = t_i = \cdots = t_s = \cdots = t_{N-2} = \frac{1}{N}$$

solve (35) for every $i = 0, 1, \dots, N-2$. The regularity of (35) can be proved by induction. For $N = 4$ the system of linear equations (35) has the form

$$\begin{aligned} t_0 \cdot 3 + t_1 \cdot 2 + t_2 \cdot 1 &= \frac{1}{2} \cdot 3, \\ t_0 \cdot 2 + t_1 \cdot 4 + t_2 \cdot 2 &= \frac{1}{2} \cdot 4, \\ t_0 \cdot 1 + t_1 \cdot 2 + t_2 \cdot 3 &= \frac{1}{2} \cdot 3, \end{aligned}$$

and it is regular, having the solution $t_0 = t_1 = t_2 = \frac{1}{4}$. $\square$

Proof of 5. To compute minimum of the mean square worst-case error (31) we start with (5) in one-dimensional form

$$\begin{aligned} & \frac{1}{N^2} \sum_{m,n=0}^{N-1} \int_0^1 \int_0^1 K(x,y) dx dy g_{m,n}(x,y) - \int_0^1 \int_0^1 K(x,y) dx dy \\ & = \frac{1}{N^2} \sum_{m=0}^{N-1} \int_0^1 \int_0^1 K(x,y) dx dy \min(x,y) \\ & + \frac{2}{N^2} \sum_{\substack{m \leq n \\ m=n=0}}^{N-1} \int_0^1 \int_0^1 K(x,y) dx dy g_{m,n}(x,y) - \int_0^1 \int_0^1 K(x,y) dx dy. \end{aligned} \quad (39)$$

We have

$$\begin{aligned} \int_0^1 \int_0^1 K(x, y) dx dy \min(x, y) &= \int_0^1 K(x, x) dx = \frac{1}{2}, \\ \int_0^1 \int_0^1 K(x, y) dx dy &= \frac{1}{3}, \end{aligned} \quad (40)$$

and

$$\begin{aligned} &\int_0^1 \int_0^1 K(x, y) dx dy g_{m,n}(x, y) \\ &= \int_0^{T_{m,n}} K(x, x + 1 - T_{m,n}) dx + \int_{T_{m,n}}^1 K(x, x - T_{m,n}) dx \\ &= \frac{T_{m,n}^2}{2} + \frac{(1 - T_{m,n})^2}{2}. \end{aligned} \quad (41)$$

Summing (41) for $T_{m,n} = \frac{n-m}{N}$, we have

$$\begin{aligned} &\sum_{\substack{m < n \\ m=n=0}}^{N-1} \left(\left(\frac{n-m}{N} \right)^2 + \left(1 - \frac{n-m}{N} \right)^2 \right) \\ &= (N-1) \left(\left(\frac{1}{N} \right)^2 + \left(1 - \frac{1}{N} \right)^2 \right) + (N-2) \left(\left(\frac{2}{N} \right)^2 + \left(1 - \frac{2}{N} \right)^2 \right) \\ &\quad + (N-3) \left(\left(\frac{3}{N} \right)^2 + \left(1 - \frac{3}{N} \right)^2 \right) + \cdots + 1 \cdot \left(\left(\frac{N-1}{N} \right)^2 + \left(1 - \frac{N-1}{N} \right)^2 \right) \\ &= \frac{1}{N} \frac{(N-1)N(2N-1)}{6}. \end{aligned} \quad (42)$$

Input (40) and (42) into (39) we find the minimum (31). $\square$

Thus the proof of Theorem 9 is finished.

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