

ON SEQUENCES INVOLVING PRIMES

YUKIO OHKUBO

ABSTRACT. In this paper we give the set of distribution functions of $f(p_n) \bmod 1$ for a special class of functions $f(x)$. Also, we generalize a result on distribution functions shown by O. Strauch and O. Blažeková to a multi-dimensional case. Applying it, we prove that every uniform distributed sequence $x_n \bmod 1$ and $\log p_n \bmod 1$ is statistically independent. By combining this and a result of G. Rauzy we derive that the sequences $p_n\theta + \log p_n$ and $p_n\theta + p_n/n$ are uniformly distributed mod1 with irrational θ .

Communicated by Vladimír Baláž

1. Introduction

For a real number x , let $[x]$ be the integral part of x , let $x \bmod 1 = \{x\} = x - [x]$ be the fractional part of x , and let $\|x\| = \min(\{x\}, 1 - \{x\})$. For $\mathbf{x} = (x_1, \dots, x_s) \in \mathbb{R}^s$, let $\mathbf{x} \bmod 1 = (x_1, \dots, x_s) \bmod 1 = (\{x_1\}, \dots, \{x_s\})$. Let c_J be the characteristic function of the interval $J \subset [0, 1]^s$.

A function $g : [0, 1]^s \rightarrow [0, 1]$ is called a distribution function (abbreviated d.f.) if $g(\mathbf{x}) = 0$ for any \mathbf{x} with a vanishing coordinate, $g(1, \dots, 1) = 1$, and $g(\mathbf{x})$ is non-decreasing, i.e.

$$\Delta_{z_s - y_s}^{(s)} \cdots \Delta_{z_1 - y_1}^{(1)} g(y_1, \dots, y_s) = \Delta_{z_s - y_s}^{(s)} (\cdots (\Delta_{z_1 - y_1}^{(1)} g(y_1, \dots, y_s))) \geq 0$$

for any $0 \leq y_i < z_i \leq 1$ ($i = 1, 2, \dots$), where $\Delta_{h_i}^{(i)} g(x_1, \dots, y_i, \dots, x_s) = g(x_1, \dots, y_i + h_i, \dots, x_s) - g(x_1, \dots, y_i, \dots, x_s)$ for $0 \leq y_i \leq y_i + h_i \leq 1$ (see [14, p.1-61]). It is easy to show that $\Delta_{z_s - y_s}^{(s)} \cdots \Delta_{z_1 - y_1}^{(1)} g(y_1, \dots, y_s) \leq 1$ for any $0 \leq y_i < z_i \leq 1$ ($i = 1, 2, \dots$), and $g(x_1, \dots, x_s)$ is non-decreasing with respect to x_i for $i = 1, \dots, s$. The d.f. $g(1, \dots, 1, x_{i_1}, 1, \dots, 1, x_{i_2}, 1, \dots, 1, x_{i_m}, 1, \dots, 1)$ is called an m -dimensional face distribution function of g in variables $(x_{i_1}, x_{i_2}, \dots, x_{i_m}) \in (0, 1)^m$, $m = 1, \dots, s - 1$.

2010 Mathematics Subject Classification: Primary 11K06; Secondary 11K36, 11K31.
Keywords: Distribution function, uniform distribution mod1, prime numbers.

We identify two d.f.s $g(\mathbf{x})$ and $\tilde{g}(\mathbf{x})$ if

- (i) $g(\mathbf{x}) = \tilde{g}(\mathbf{x})$ at every common point $\mathbf{x} \in (0, 1)^s$ of continuity, and
- (ii) $g(1, \dots, 1, x_{i_1}, 1, \dots, 1, x_{i_m}, 1, \dots, 1) = \tilde{g}(1, \dots, 1, x_{i_1}, 1, \dots, 1, x_{i_m}, 1, \dots, 1)$ for every common point $(x_{i_1}, \dots, x_{i_m}) \in (0, 1)^m$ of continuity in every m -dimensional face distribution function of g and \tilde{g} , $m = 1, \dots, s - 1$ (see [14, p.1-61]).

For two s -dimensional sequences of real numbers $\mathbf{a}_n = (a_{n,1}, \dots, a_{n,s})$, $n = 1, 2, \dots$ and $\tilde{\mathbf{a}}_n = (\tilde{a}_{n,1}, \dots, \tilde{a}_{n,s})$, $n = 1, 2, \dots$, we define the empirical distribution by

$$F_N(\mathbf{x}) = \frac{1}{N} \sum_{n=1}^N c_{[0,x_1] \times \dots \times [0,x_s]}(\{a_{n,1}\}, \dots, \{a_{n,s}\}),$$

$$\tilde{F}_N(\mathbf{x}) = \frac{1}{N} \sum_{n=1}^N c_{[0,x_1] \times \dots \times [0,x_s]}(\{\tilde{a}_{n,1}\}, \dots, \{\tilde{a}_{n,s}\}),$$

for all $\mathbf{x} = (x_1, \dots, x_s) \in [0, 1]^s$.

If there exists an increasing sequence of positive integers N_k , $k = 1, 2, \dots$, such that

$$\lim_{k \rightarrow \infty} F_{N_k}(\mathbf{x}) = g(\mathbf{x})$$

at every continuity point $\mathbf{x} \in (0, 1)^s$ of $g(\mathbf{x})$, and

$$\lim_{k \rightarrow \infty} F_{N_k}(1, \dots, 1, x_{i_1}, 1, \dots, 1, x_{i_m}, 1, \dots, 1) = g(1, \dots, 1, x_{i_1}, 1, \dots, 1, x_{i_m}, 1, \dots, 1)$$

at every continuity point $(x_{i_1}, \dots, x_{i_m}) \in [0, 1]^m$ of $g(1, \dots, 1, x_{i_1}, 1, \dots, 1, x_{i_m}, 1, \dots, 1)$ for $m = 1, \dots, s - 1$, then the function $g(\mathbf{x})$ is called the distribution function (abbreviated d.f.) of the sequence $\mathbf{a}_n \bmod 1$, $n = 1, 2, \dots$. We denote as $G(\mathbf{a}_n \bmod 1)$ the set of all d.f.s of the sequence $\mathbf{a}_n \bmod 1$, $n = 1, 2, \dots$ (see [14, p.1-61]). The sequence \mathbf{a}_n is said to be uniformly distributed mod 1 if

$$\lim_{N \rightarrow \infty} F_N(x_1, \dots, x_s) = x_1 \cdots x_s$$

for all $(x_1, \dots, x_s) \in [0, 1]^s$.

In [12] O. Strauch and O. Blažeková showed the following theorem.

THEOREM A ([12], Theorem 1). *Let x_n and y_n be two real sequences. Assume that all d.f.s in $G(x_n \bmod 1)$ are continuous at 0 and 1. Then the zero limit of fractional parts*

$$\lim_{n \rightarrow \infty} (x_n - y_n) \bmod 1 = 0$$

implies $G(x_n \bmod 1) = G(y_n \bmod 1)$.

R. Giuliano Antonini and O. Strauch [3] extended this theorem to the case of weighted distribution functions. In [3], the referee notes that if $x_n, y_n \in [0, 1)$, $n = 1, 2, \dots$, then the assumption that all d.f.s in $G(x_n \bmod 1)$ are continuous at 0 and 1 can be omitted.

Let p_n denote the n th prime. Applying Theorem A, O. Strauch and O. Blažeková [12] proved that

$$G(p_n/n \bmod 1) = G(\log(n \log n) \bmod 1).$$

Furthermore, by two methods they proved

$$G(\log(n \log n) \bmod 1) = G(\log n \bmod 1). \tag{1.1}$$

It is well-known that

$$G(\log n \bmod 1) = \left\{ g_u(x) = \frac{e^x - 1}{e - 1} e^{-u} + (e^{\min(x,u)} - 1)e^{-u} : u \in [0, 1] \right\} \tag{1.2}$$

(see [6, pp.57-58], [14, 2.12]).

This paper is divided into three parts. The first part (Section 2) gives the set of d.f.s of $f(p_n) \bmod 1$ for a special class of functions $f(x)$ that is described in O. Strauch and O. Blažeková [12, Theorem 5].

In the second part (Section 3) we give a generalization of Theorem A to a multidimensional case. Applying it, we prove that every uniform distributed sequence $x_n \bmod 1$ and $\log p_n \bmod 1$ is statistically independent. By combining this and a result of G. Rauzy [8] (see Lemma B), we derive that the sequences $p_n \theta + \log p_n$ and $p_n \theta + p_n/n$ are uniformly distributed mod 1.

In the third part (Section 4) we give the proof of Theorem 2.

2. Distribution of $f(p_n) \bmod 1$

Hadamard and de la Vallée Poussin proved the prime number theorem

$$\pi(x) = \text{Li}(x) + O(xe^{-A\sqrt{\log x}}),$$

for some positive constant A (cf. P. Ribenboim [9, p.226]). By integration by parts

$$\text{Li}(x) = \frac{x}{\log x - 1} - \frac{x}{(\log x)^2(\log x - 1)} - \frac{2}{\log 2} - \frac{2}{(\log 2)^2} + 2 \int_2^x \frac{1}{(\log t)^3} dt.$$

Since

$$\int_2^x \frac{1}{(\log t)^3} dt = \int_2^{x^{1/2}} \frac{1}{(\log t)^3} dt + \int_{x^{1/2}}^x \frac{1}{(\log t)^3} dt \ll \frac{x}{(\log x)^3},$$

we have

$$\pi(x) = \frac{x}{\log x - 1} + O\left(\frac{x}{(\log x)^3}\right).$$

Hence

$$\frac{p_n}{n} = \log p_n - 1 + O\left(\frac{1}{\log p_n}\right). \tag{2.1}$$

On the other hand, it is known that

$$\frac{p_n}{n} = \log n + (\log \log n - 1) + o\left(\frac{\log \log n}{\log n}\right) \tag{2.2}$$

(see [9], [12]). From (2.1) and (2.2) it follows that

$$\log p_n = \log(n \log n) + o\left(\frac{\log \log n}{\log n}\right) + O\left(\frac{1}{\log p_n}\right).$$

Therefore we have

$$\lim_{n \rightarrow \infty} \|\log p_n - \log(n \log n)\| = 0. \tag{2.3}$$

Applying Theorem A with the continuity of $g(x) \in G(\log(n \log n) \bmod 1)$ at $x = 0$ and $x = 1$, we have

$$G(\log p_n \bmod 1) = G(\log(n \log n) \bmod 1).$$

Since $G(\log(n \log n) \bmod 1) = G(\log n \bmod 1)$ by (1.1),

$$G(\log p_n \bmod 1) = G(\log n \bmod 1).$$

The d.f.s of the sequence $\log p_n \bmod 1$ was studied by A. Wintner [16]. For a function $f(x)$ that satisfies conditions of the following theorem, O. Strauch and O. Blažeková showed some extension of the Koksma theorem [7] about the sequence $f(n)$ (see [12, Theorem 5]). They applied it to $f(x) = \log x$ and $f(x) = \log(x \log^{(i)} x)$, $i = 1, 2, \dots$, where $\log^{(i)} x$ is the i th iterated logarithm $\log \cdots \log x$. We prove the following result for the sequence $f(p_n) \bmod 1$ by using the methods of [16] and [12, Theorem 5].

THEOREM 1. *Let the real-valued function $f(x)$ be strictly increasing for $x \geq 1$ and let $f^{-1}(x)$ be the inverse function of $f(x)$. Suppose that*

- (i) $\lim_{k \rightarrow \infty} f^{-1}(k+1) - f^{-1}(k) = \infty$,
- (ii) $\lim_{k \rightarrow \infty} \frac{f^{-1}(k+w_k)}{f^{-1}(k)} = \psi(u)$ for every sequence $w_k \in [0, 1]$

ON SEQUENCES INVOLVING PRIMES

for which $\lim_{k \rightarrow \infty} w_k = u$, where this limit defines the function $\psi(u)$ on $[0, 1]$,
 (iii) $\psi(1) > 1$. Then

$$G(f(p_n) \bmod 1) = \left\{ \tilde{g}_u(x) = \frac{\min(\psi(x), \psi(u)) - 1}{\psi(u)} + \frac{1}{\psi(u)} \cdot \frac{\psi(x) - 1}{\psi(1) - 1} : u \in [0, 1] \right\}.$$

Proof. First we show that the sequence $\{f(p_n)\}$ is dense in $[0, 1]$. For $0 \leq a < b \leq 1$, $a < \{f(p_n)\} \leq b$ if and only if there exists $k \in \mathbb{Z}$ such that $f^{-1}(k+a) < p_n \leq f^{-1}(k+b)$. For $\xi > 0$, let $\pi(\xi)$ be the number of primes p with $p \leq \xi$. By the prime number theorem, we have

$$\begin{aligned} & \pi(f^{-1}(k+b)) - \pi(f^{-1}(k+a)) \\ &= \frac{f^{-1}(k+b)}{\log f^{-1}(k+b)} - \frac{f^{-1}(k+a)}{\log f^{-1}(k+a)} + O\left(\frac{f^{-1}(k+b)}{\log^2 f^{-1}(k+b)}\right) \\ &= \frac{f^{-1}(k+b)}{\log f^{-1}(k+b)} \left(1 - \frac{f^{-1}(k+a)}{f^{-1}(k+b)} \frac{\log f^{-1}(k+b)}{\log f^{-1}(k+a)}\right) + O\left(\frac{f^{-1}(k+b)}{\log^2 f^{-1}(k+b)}\right). \end{aligned}$$

From (i) it follows that $\lim_{x \rightarrow \infty} f^{-1}(x) = \infty$, thus $\lim_{k \rightarrow \infty} \frac{f^{-1}(k+b)}{\log f^{-1}(k+b)} = \infty$. (ii) implies $\lim_{k \rightarrow \infty} \frac{f^{-1}(k+a)}{f^{-1}(k+b)} = \frac{\psi(a)}{\psi(b)} < 1$ and

$$\lim_{k \rightarrow \infty} \frac{\log f^{-1}(k+b)}{\log f^{-1}(k+a)} = \lim_{k \rightarrow \infty} \left(\frac{1}{\log f^{-1}(k+a)} \log \frac{f^{-1}(k+b)}{f^{-1}(k+a)} \right) + 1 = 1.$$

Therefore, $\pi(f^{-1}(k+b)) - \pi(f^{-1}(k+a)) \rightarrow \infty$ as $k \rightarrow \infty$. Hence, there exist k and n such that $f^{-1}(k+a) < p_n \leq f^{-1}(k+b)$. So it was shown that the sequence $\{f(p_n)\}$ is dense in $[0, 1]$.

Since the sequence $\{f(p_n)\}$ is dense in $[0, 1]$, for $0 \leq u \leq 1$ there exists a sequence of positive integers n_j such that $u_j = \{f(p_{n_j})\} \rightarrow u$ as $j \rightarrow \infty$. For $\xi > 0$, put $N(\xi) = \#\{n \geq 1 | f(p_n) < \xi\}$. For $0 \leq x \leq 1$ and $m \geq 1$, put $\phi(m; x) = \frac{1}{m} \#\{k \leq m | \{f(p_k)\} < x\}$. Then we have

$$N(\xi)\phi(N(\xi); x) = \sum_{k=0}^{[\xi]-1} (N(k+x) - N(k)) + N(\min(\xi, [\xi] + x)) - N([\xi]).$$

The prime number theorem implies

$$N(\xi) = \#\{n | p_n < f^{-1}(\xi)\} = \pi(f^{-1}(\xi)) + c(\xi) \sim \frac{f^{-1}(\xi)}{\log f^{-1}(\xi)}, \quad (2.4)$$

where $c(\xi) = 0$ or 1 .

Hence, we have

$$\begin{aligned} \phi(n_j - 1; x) &= \phi(N(f(p_{n_j})); x) \\ &= \frac{1}{N(f(p_{n_j}))} \sum_{k=0}^{h_j-1} (N(k+x) - N(k)) + \frac{N(h_j + \min(u_j, x))}{N(f(p_{n_j}))} - \frac{N(h_j)}{N(f(p_{n_j}))} \\ &= \frac{\sum_{k=0}^{h_j-1} (N(k+x) - N(k))}{\sum_{k=0}^{h_j-1} (N(k+1) - N(k))} \cdot \frac{N(h_j) - N(0)}{N(h_j + u_j)} \\ &\quad + \frac{N(h_j + \min(u_j, x))}{N(h_j + u_j)} - \frac{N(h_j)}{N(h_j + u_j)}, \end{aligned}$$

where $h_j = [f(p_{n_j})]$. Let $0 \leq x \leq 1$ be fixed. Since $N(n)$ is strictly increasing and tends to infinity, Cauchy-Stolz lemma implies

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{\sum_{k=0}^n (N(k+x) - N(k))}{\sum_{k=0}^n (N(k+1) - N(k))} &= \lim_{k \rightarrow \infty} \frac{N(k+x) - N(k)}{N(k+1) - N(k)} \\ &= \lim_{k \rightarrow \infty} \frac{N(k+x)/N(k) - 1}{N(k+1)/N(k) - 1}. \end{aligned}$$

From (2.4), we obtain

$$\frac{N(k+x)}{N(k)} \sim \frac{\log f^{-1}(k)}{\log f^{-1}(k+x)} \cdot \frac{f^{-1}(k+x)}{f^{-1}(k)}.$$

Furthermore, (i) and (ii) yield

$$\lim_{k \rightarrow \infty} \frac{\log f^{-1}(k)}{\log f^{-1}(k+x)} = \lim_{k \rightarrow \infty} \left(\frac{1}{\log f^{-1}(k+x)} \log \frac{f^{-1}(k)}{f^{-1}(k+x)} \right) + 1 = 1.$$

Thus, we get

$$\lim_{k \rightarrow \infty} \frac{N(k+x)}{N(k)} = \psi(x),$$

and so,

$$\lim_{j \rightarrow \infty} \frac{\sum_{k=0}^{h_j-1} (N(k+x) - N(k))}{\sum_{k=0}^{h_j-1} (N(k+1) - N(k))} = \lim_{n \rightarrow \infty} \frac{\sum_{k=0}^n (N(k+x) - N(k))}{\sum_{k=0}^n (N(k+1) - N(k))} = \frac{\psi(x) - 1}{\psi(1) - 1}.$$

Using (ii) with $w_{h_j} = u_j$, $w_k = u$ otherwise, we obtain

$$\lim_{j \rightarrow \infty} \frac{f^{-1}(h_j + u_j)}{f^{-1}(h_j)} = \psi(u), \quad \lim_{j \rightarrow \infty} \frac{\log f^{-1}(h_j)}{\log f^{-1}(h_j + u_j)} = 1.$$

and hence

$$\lim_{j \rightarrow \infty} \frac{N(h_j)}{N(h_j + u_j)} = \lim_{j \rightarrow \infty} \frac{\log f^{-1}(h_j + u_j)}{\log f^{-1}(h_j)} \cdot \frac{f^{-1}(h_j)}{f^{-1}(h_j + u_j)} = \frac{1}{\psi(u)}.$$

Let $u < x < 1$ be fixed. Since $\min(u_j, x) = u_j$ for sufficiently large j , we obtain

$$\lim_{j \rightarrow \infty} \frac{N(h_j + \min(u_j, x))}{N(h_j + u_j)} = 1.$$

Let $0 \leq x < u$ be fixed. For sufficiently large j , we have

$$\frac{N(h_j + \min(u_j, x))}{N(h_j + u_j)} = \frac{N(h_j + x)}{N(h_j + u_j)} \sim \frac{f^{-1}(h_j + x)/f^{-1}(h_j)}{f^{-1}(h_j + u_j)/f^{-1}(h_j)} \cdot \frac{\log f^{-1}(h_j + u_j)}{\log f^{-1}(h_j + x)}.$$

From (i) and (ii) it follows that

$$\lim_{j \rightarrow \infty} \frac{\log f^{-1}(h_j + u_j)}{\log f^{-1}(h_j + x)} = \lim_{j \rightarrow \infty} \left(\frac{1}{\log f^{-1}(h_j + x)} \log \frac{f^{-1}(h_j + u_j)}{f^{-1}(h_j + x)} \right) + 1 = 1.$$

Thus, we get

$$\lim_{j \rightarrow \infty} \frac{N(h_j + \min(u_j, x))}{N(h_j + u_j)} = \frac{\psi(x)}{\psi(u)}.$$

Therefore,

$$\lim_{j \rightarrow \infty} \frac{N(h_j + \min(u_j, x))}{N(h_j + u_j)} = \frac{\min(\psi(x), \psi(u))}{\psi(u)}.$$

Consequently,

$$\tilde{g}_u(x) = \lim_{j \rightarrow \infty} \phi(n_j - 1; x) = \frac{\min(\psi(x), \psi(u)) - 1}{\psi(u)} + \frac{1}{\psi(u)} \cdot \frac{\psi(x) - 1}{\psi(1) - 1}.$$

Conversely, if $g(x) \in G(f(p_n) \bmod 1)$, then there exists a sequence m_i such that $\lim_{i \rightarrow \infty} \phi(m_i; x) = g(x)$. The sequence $m_i + 1$ contains a subsequence n_j such that $\lim_{j \rightarrow \infty} \{f(p_{n_j})\} = u$ for some $0 \leq u \leq 1$. Therefore, we have $g(x) = \tilde{g}_u(x)$. \square

As O. Strauch and O. Blažeková [12] applied Theorem 5 of [12] to $f(x) = \log x$ and $f(x) = \log(x \log^{(i)} x)$, $i = 1, 2, \dots$, we can apply Theorem 1 to the same functions. In consequence, it is shown that the sequence $\log p_n$ and the sequences $\log(p_n \log^{(i)} p_n)$, $i = 1, 2, \dots$ have the same distribution as one of the sequence $\log n$. This solves the Open problem 1.3 in [13]. S. Akiyama [1] proved that the sequence $\sum_{i=0}^{\ell-1} c_i \log p_{n+i}$ is not almost uniformly distributed modulo 1, where

$c_i, i = 0, 1, 2, \dots$, are real numbers with $\sum_{i=0}^{\ell-1} c_i \neq 0$, i.e. $x \notin G(\sum_{i=0}^{\ell-1} c_i \log p_{n+i})$ (see also [14, 2.19.8]). Akiyama [1] used the following estimation

$$\sum_{i=0}^{\ell-1} c_i \log p_{n+i} - \left(\sum_{i=0}^{\ell-1} c_i \right) \log p_n = o(1).$$

Therefore by Theorem A (or Theorem 2),

$$G \left(\sum_{i=0}^{\ell-1} c_i \log p_{n+i} \bmod 1 \right) = G \left(\left(\sum_{i=0}^{\ell-1} c_i \right) \log p_n \bmod 1 \right).$$

On the other hand, by Theorem 1

$$G(c \log p_n \bmod 1) = \left\{ \frac{e^{x/c} - 1}{e^{1/c} - 1} e^{-u/c} + (e^{\min(x/c, u/c)} - 1) e^{-u/c} : u \in [0, 1] \right\},$$

where $c = \sum_{i=0}^{\ell-1} c_i$. Hence, the sequence $\sum_{i=0}^{\ell-1} c_i \log p_{n+i}$ has the same distribution.

3. Distribution of the another type of sequence

We extend Theorem A to a multidimensional case as follows.

THEOREM 2. *Let $\mathbf{a}_n = (a_{n,1}, \dots, a_{n,s})$, $\tilde{\mathbf{a}}_n = (\tilde{a}_{n,1}, \dots, \tilde{a}_{n,s})$, $n = 1, 2, \dots$ be s -dimensional sequences of real numbers. Suppose that for every d.f. $g(x_1, \dots, x_s) \in G(\mathbf{a}_n \bmod 1)$, the face d.f. $g(1, \dots, 1, x_i, 1, \dots)$ is continuous at $x_i = 0, 1$ for $i = 1, \dots, s$. If $\|a_{n,i} - \tilde{a}_{n,i}\| \rightarrow 0$ ($n \rightarrow \infty$) for $i = 1, \dots, s$, then $G(\mathbf{a}_n \bmod 1) = G(\tilde{\mathbf{a}}_n \bmod 1)$.*

The proof is described in Section 4. We recall definitions and results on statistical independence.

DEFINITION 1 (see [8, p.91]). *Let x_n and y_n be two infinite real sequences. Two sequences $x_n \bmod 1$ and $y_n \bmod 1$ are called statistically independent if*

$$\lim_{N \rightarrow \infty} \left(\frac{1}{N} \sum_{n=1}^N f(\{x_n\}) g(\{y_n\}) - \left(\sum_{n=1}^N f(\{x_n\}) \right) \left(\sum_{n=1}^N g(\{y_n\}) \right) \right) = 0$$

for every continuous real-valued functions f, g defined on $[0, 1]$.

ON SEQUENCES INVOLVING PRIMES

LEMMA A (see [8, pp.97-98]). *Let x_n and y_n be two infinite real sequences. Two sequences $x_n \bmod 1$ and $y_n \bmod 1$ are statistically independent if and only if*

$$\lim_{N \rightarrow \infty} \left(\frac{1}{N} \sum_{n=1}^N e^{2\pi i(hx_n + ky_n)} - \left(\frac{1}{N} \sum_{n=1}^N e^{2\pi ihx_n} \right) \left(\frac{1}{N} \sum_{n=1}^N e^{2\pi iky_n} \right) \right) = 0$$

for every integers $h \neq 0, k \neq 0$.

LEMMA 1. *Let $x_n, n = 1, 2, 3 \dots$ be a real sequence. If x_n is uniformly distributed mod 1, then $x_n \bmod 1$ and $\log(n \log n) \bmod 1$ are statistically independent.*

Proof. Suppose that x_n is uniformly distributed mod 1. By Weyl criterion we have

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N e^{2\pi ihx_n} = 0. \tag{3.1}$$

Therefore by applying Lemma A it is sufficient to show that

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N e^{2\pi i(hx_n + k \log(n \log n))} = 0. \tag{3.2}$$

By Abel summation formula we obtain

$$\begin{aligned} & \sum_{n=1}^N e^{2\pi i(hx_n + k \log(n \log n))} \\ = & \sum_{n=1}^{N-1} \left(e^{2\pi ik \log(n \log n)} - e^{2\pi ik \log((n+1) \log(n+1))} \right) \sum_{m=1}^n e^{2\pi ihx_m} \\ & + e^{2\pi ik \log(N \log N)} \sum_{n=1}^N e^{2\pi ihx_n}. \end{aligned}$$

From

$$\left| e^{2\pi ik \log(n \log n)} - e^{2\pi ik \log((n+1) \log(n+1))} \right| \leq 2\pi |k| \frac{\log n + 1}{n \log n}$$

and (3.1) we obtain (3.2). □

THEOREM B (see [8, p.92], [4, Theorem 2]). *Let x_n and y_n be two infinite real sequences. Two sequences $x_n \bmod 1$ and $y_n \bmod 1$ are statistically independent if and only if for every $g \in G((x_n, y_n) \bmod 1)$*

$$g(x, y) = g(x, 1)g(1, y) \quad \text{a.e. on } [0, 1]^2.$$

LEMMA 2. *Let $x_n, n = 1, 2, \dots$ be a real sequence. If x_n is uniformly distributed mod 1, then $G((x_n, \log p_n) \bmod 1) = G((x_n, \log(n \log n)) \bmod 1) = \{xg_u(y) : u \in [0, 1]\}$.*

Proof. Let $g \in G((x_n, \log(n \log n)) \bmod 1)$. From Lemma 1, it follows that $x_n \bmod 1$ and $\log(n \log n) \bmod 1$ are statistically independent. Applying Theorem B, we have

$$g(x, y) = g(x, 1)g(1, y) \text{ a.e. on } [0, 1]^2$$

Since $g(x, 1) \in G(x_n \bmod 1) = \{x\}$, we have $g(x, 1) = x$ at every continuity point x of $g(x, 1)$. We have $g(x, 1) = x$ for all $x \in [0, 1]$. Since $g(1, y) \in G(\log(n \log n) \bmod 1)$, by (1.2) there exists $u \in [0, 1]$ such that $g(1, y) = g_u(y)$ at every continuity point y of $g(1, y)$. Since $g_u(y)$ is continuous, we obtain $g(1, y) = f_u(y)$ for every $y \in [0, 1]$. It follows that

$$g(x, y) = xg_u(y) \text{ for every } (x, y) \in [0, 1]^2.$$

By this and (2.3), Theorem 2 yields

$$G((x_n, \log p_n) \bmod 1) = G((x_n, \log(n \log n)) \bmod 1). \quad \square$$

THEOREM 3. *Let $x_n, n = 1, 2, \dots$ be a real sequence. If x_n is uniformly distributed mod 1, then $x_n \bmod 1$ and $\log p_n \bmod 1$ is statistically independent.*

Proof. Let $g \in G((x_n, \log p_n) \bmod 1)$. By Lemma 2 there exist $\tilde{g} \in G((x_n, \log(n \log n)) \bmod 1)$ such that $g(x, y) = \tilde{g}(x, y)$ at every continuity point $(x, y) \in [0, 1]^2$. Since $\tilde{g}(x, y)$ is continuous on $[0, 1]^2$,

$$g(x, y) = \tilde{g}(x, y) \tag{3.3}$$

for every $(x, y) \in [0, 1]^2$. By Lemma 1 and Theorem B we have $\tilde{g}(x, y) = \tilde{g}(x, 1)\tilde{g}(1, y)$ almost everywhere. Hence from (3.3) it follows that

$$g(x, y) = g(x, 1)g(1, y)$$

almost everywhere. This and Theorem B complete the proof of Theorem 3. \square

LEMMA B (see [8, p.96], [14]). *Let x_n and y_n be sequences of real numbers such that $x_n \bmod 1$ and $y_n \bmod 1$ are statistically independent, and x_n is uniformly distributed mod 1. Then the sequence $x_n + y_n$ is uniformly distributed mod 1.*

It is well known that $p_n\theta$ is uniformly distributed mod 1, for an irrational number θ (see [15, Chapter 11]). By this result, Theorem 3 and Lemma B we obtain the following result.

THEOREM 4. *If θ is irrational, then $p_n\theta + \log p_n$ is uniformly distributed mod 1.*

Furthermore, we have the following result in same way as the proof of Theorem 4. If θ is irrational, then $p_n(\theta + 1/n)$ is uniformly distributed mod 1.

4. Proof of Theorem 2

In order to prove Theorem 2 we need some properties of monotone functions of several variables in [2].

DEFINITION 2 (see [2]). *Let $F(x_1, \dots, x_s)$ be a real valued function defined on the interval $J = [0, 1]^s$. The function F is said to be monotone if $F(x_1, \dots, x_s)$ is non-decreasing with respect to x_i for $1 \leq i \leq s$, and if $\Delta_{h_k}^{(k)}(\Delta_{h_j}^{(j)} F(x_1, \dots, x_s)) \geq 0$ for $0 \leq x_i \leq 1$ ($i = 1, \dots, s$), $1 \leq j < k \leq s$, $h_j \geq 0$, $h_k \geq 0$, $x_j + h_j \leq 1$, $x_k + h_k \leq 1$.*

LEMMA C (see [2], Theorem 1). *If $F(x_1, \dots, x_s)$ be a monotone function on the interval J , then there is a set R consisting of the points of a countable collection $(s - 1)$ -space $x_i = \alpha_i^{(j)}$, $i = 1, 2, \dots, s$, $j = 1, 2, \dots$, such that F is continuous at all points of $J - R$.*

REMARK 1. *Since $F(1, \dots, 1, x_i, 1, \dots, 1)$ is non-decreasing with respect to x_i , it has a countable number of discontinuities $x_i = \alpha_{ij}$, $i = 1, \dots, s$, $j = 1, 2, \dots$. The set R is the union of the sets in $[0, 1]^s$: $x_i = \alpha_{ij}$, $i = 1, \dots, s$, $j = 1, 2, \dots$*

LEMMA D (see [2], Theorem 2). *Let $F_n(x_1, \dots, x_s)$ be a sequence of monotone functions on J with $|F_n(x_1, \dots, x_s)| \leq A$ ($n = 1, 2, \dots$) for some $A > 0$. Then, there exists a sequence of positive integers $n_1 < n_2 < \dots$ and a non-decreasing function $F(x_1, \dots, x_s)$ such that*

$$\lim_{k \rightarrow \infty} F_{n_k}(x_1, x_2, \dots, x_s) = F(x_1, x_2, \dots, x_s) \quad \text{on } J.$$

LEMMA 3. *Let g be a distribution function. If there exist y_{i,n_i} and z_{i,n_i} ($n_i = 1, 2, \dots$) for $i = 1, 2, \dots, s$ such that $0 < y_{i,n_i} < z_{i,n_i} < 1$ ($i = 1, 2, \dots, s$), $\lim_{n_i \rightarrow \infty} y_{i,n_i} = 0$, $\lim_{n_i \rightarrow \infty} z_{i,n_i} = 1$, and*

$$\begin{aligned} \lim_{\substack{n_i \rightarrow \infty \\ i=1 \dots s}} \Delta_{z_{s,n_s}-y_{s,n_s}}^{(s)} \dots \Delta_{z_{1,n_1}-y_{1,n_1}}^{(1)} g(y_{1,n_1}, \dots, y_{s,n_s}) &:= \\ \lim_{n_s \rightarrow \infty} \dots \lim_{n_1 \rightarrow \infty} \Delta_{z_{s,n_s}-y_{s,n_s}}^{(s)} \dots \Delta_{z_{1,n_1}-y_{1,n_1}}^{(1)} g(y_{1,n_1}, \dots, y_{s,n_s}) &= 1. \end{aligned}$$

then

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1 \dots s}} g(z_{1,n_1}, \dots, z_{s,n_s}) = 1 \quad \text{and} \quad \lim_{\substack{n_i \rightarrow \infty \\ i=1 \dots s}} g(t_{1,n_1}, \dots, t_{s,n_s}) = 0,$$

where $t_{j,n_j} = y_{j,n_j}$ for some j and either $t_{i,n_i} = y_{i,n_i}$ or $t_{i,n_i} = z_{i,n_i}$ for $i = 1, \dots, s$.

Proof. By induction on s we shall prove this lemma.

For $s = 1$, from $\lim_{n_1 \rightarrow \infty} (g(z_{1,n_1}) - g(y_{1,n_1})) = 1$ and $0 \leq g(y_{1,n_1}) \leq g(z_{1,n_1}) \leq 1$ we have

$$\lim_{n_1 \rightarrow \infty} g(z_{1,n_1}) = 1, \quad \lim_{n_1 \rightarrow \infty} g(y_{1,n_1}) = 0.$$

For $s > 1$ we have

$$\begin{aligned} & \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s}} \Delta_{z_{s,n_s} - y_{s,n_s}}^{(s)} \cdots \Delta_{z_{1,n_1} - y_{1,n_1}}^{(1)} g(y_{1,n_1}, \cdots, y_{s,n_s}) \\ &= \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \Delta_{z_{s-1,n_{s-1}} - y_{s-1,n_{s-1}}}^{(s-1)} \cdots \Delta_{z_{1,n_1} - y_{1,n_1}}^{(1)} \\ & \quad \lim_{n_s \rightarrow \infty} g(y_{1,n_1}, \cdots, y_{s-1,n_{s-1}}, z_{s,n_s}) - \\ & \quad \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \Delta_{z_{s-1,n_{s-1}} - y_{s-1,n_{s-1}}}^{(s-1)} \cdots \Delta_{z_{1,n_1} - y_{1,n_1}}^{(1)} \\ & \quad \lim_{n_s \rightarrow \infty} g(y_{1,n_1}, \cdots, y_{s-1,n_{s-1}}, y_{s,n_s}) = 1. \end{aligned}$$

Set

$$\begin{aligned} g_0(y_1, \cdots, y_{s-1}) &:= \lim_{n_s \rightarrow \infty} g(y_1, \cdots, y_{s-1}, y_{s,n_s}), \\ g_1(y_1, \cdots, y_{s-1}) &:= \lim_{n_s \rightarrow \infty} g(y_1, \cdots, y_{s-1}, z_{s,n_s}). \end{aligned}$$

It is shown that $g_k(y_1, \cdots, y_{s-1})$ is a d.f. for $k = 0, 1$. Hence we obtain

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \Delta_{z_{s-1,n_{s-1}} - y_{s-1,n_{s-1}}}^{(s-1)} \cdots \Delta_{z_{1,n_1} - y_{1,n_1}}^{(1)} g_1(y_1, \cdots, y_{s-1,n_{s-1}}) = 1.$$

From this and the induction hypothesis we obtain

$$\begin{aligned} \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \lim_{n_s \rightarrow \infty} g(z_{1,n_1}, \cdots, z_{s-1,n_{s-1}}, z_{s,n_s}) &= \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s}} g_1(z_{1,n_1}, \cdots, z_{s-1,n_{s-1}}) = 1, \\ \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \lim_{n_s \rightarrow \infty} g(t_{1,n_1}, \cdots, t_{s-1,n_{s-1}}, z_{s,n_s}) &= \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s}} g_1(t_{1,n_1}, \cdots, t_{s-1,n_{s-1}}) = 0, \end{aligned}$$

where either $t_{i,n_i} = y_{i,n_i}$ or $t_{i,n_i} = z_{i,n_i}$ for $1 \leq i \leq s-1$ and $t_{j,n_j} = y_{j,n_j}$ for some j . Since $g(x_1, \dots, x_{s-1}, x_s)$ is non-decreasing with respect to x_s ,

$$\begin{aligned} & \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \lim_{n_s \rightarrow \infty} g(t_{1,n_1}, \cdots, t_{s-1,n_{s-1}}, y_{s,n_s}) \\ & \leq \lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s-1}} \lim_{n_s \rightarrow \infty} g(t_{1,n_1}, \cdots, t_{s-1,n_{s-1}}, z_{s,n_s}) = 0. \end{aligned}$$

Thus

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1 \cdots s}} g(t_{1,n_1}, \cdots, t_{s-1,n_{s-1}}, y_{s,n_s}) = 0.$$

□

ON SEQUENCES INVOLVING PRIMES

We prove the two dimensional case of Theorem 2. The proof is analogous to the proof of Theorem 1 in [12].

Since either $\|x_n - \tilde{x}_n\| = \{x_n - \tilde{x}_n\} = (x_n - [x_n - \tilde{x}_n]) - \tilde{x}_n$ or $\|x_n - \tilde{x}_n\| = 1 - \{x_n - \tilde{x}_n\} = \tilde{x}_n - (x_n - [x_n - \tilde{x}_n] - 1)$, we can assume that $|a_{n,i} - \tilde{a}_{n,i}| \rightarrow 0$ for $i = 1, 2$.

Let $g(\mathbf{x}) \in G(\mathbf{a}_n \bmod 1)$. There exists the sequence N'_k such that

$$\lim_{k \rightarrow \infty} F_{N'_k}(\mathbf{x}) = g(\mathbf{x})$$

at every continuity point $\mathbf{x} \in (0, 1)^2$ of $g(\mathbf{x})$.

By Lemma D there exists the subsequence $N'_{k'_j}$ of the sequence N'_k and $\tilde{g}(\mathbf{x}) \in G(\tilde{\mathbf{x}}_n \bmod 1)$ such that

$$\lim_{j \rightarrow \infty} \tilde{F}_{N'_{k'_j}}(\mathbf{x}) = \tilde{g}(\mathbf{x})$$

for all $\mathbf{x} \in [0, 1]^2$.

Thus there exists N_j such that

$$\lim_{j \rightarrow \infty} F_{N_j}(\mathbf{x}) = g(\mathbf{x})$$

at every continuity point $\mathbf{x} \in (0, 1)^2$ of $g(\mathbf{x})$, and

$$\lim_{j \rightarrow \infty} \tilde{F}_{N_j}(\mathbf{x}) = \tilde{g}(\mathbf{x})$$

for all $\mathbf{x} \in [0, 1]^2$.

Applying the second Helly theorem, we have

$$\lim_{j \rightarrow \infty} \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} dF_{N_j}(x_1, x_2) = \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} dg(x_1, x_2),$$

$$\lim_{j \rightarrow \infty} \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} d\tilde{F}_{N_j}(x_1, x_2) = \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} d\tilde{g}(x_1, x_2)$$

for $h_i = 0, \pm 1, \pm 2, \dots$ ($i = 1, 2$).

Hence

$$\lim_{j \rightarrow \infty} \frac{1}{N_j} \sum_{n=1}^{N_j} e^{2\pi i(h_1 a_{n,1} + h_2 a_{n,2})} = \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} dg(x_1, x_2),$$

$$\lim_{j \rightarrow \infty} \frac{1}{N_j} \sum_{n=1}^{N_j} e^{2\pi i(h_1 \tilde{a}_{n,1} + h_2 \tilde{a}_{n,2})} = \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} d\tilde{g}(x_1, x_2)$$

for $h_i = 0, \pm 1, \pm 2, \dots$ ($i = 1, 2$).

Since

$$\begin{aligned} & \left| \frac{1}{N_j} \sum_{n=1}^{N_j} e^{2\pi i(h_1 a_{n,1} + h_2 a_{n,2})} - \sum_{n=1}^{N_j} e^{2\pi i(h_1 \tilde{a}_{n,1} + h_2 \tilde{a}_{n,2})} \right| \\ & \leq \frac{1}{N_j} \sum_{n=1}^{N_j} |e^{2\pi i(h_1(a_{n,1} - \tilde{a}_{n,1}) + h_2(a_{n,2} - \tilde{a}_{n,2}))} - 1| \rightarrow 0 \end{aligned}$$

as $j \rightarrow \infty$, we have

$$\lim_{j \rightarrow \infty} \frac{1}{N_j} \sum_{n=1}^{N_j} e^{2\pi i(h_1 a_{n,1} + h_2 a_{n,2})} = \lim_{j \rightarrow \infty} \frac{1}{N_j} \sum_{n=1}^{N_j} e^{2\pi i(h_1 \tilde{a}_{n,1} + h_2 \tilde{a}_{n,2})}.$$

Therefore we obtain

$$\int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} dg(x_1, x_2) = \int_{[0,1]^2} e^{2\pi i(h_1 x_1 + h_2 x_2)} d\tilde{g}(x_1, x_2)$$

for $h_i = 0, \pm 1, \pm 2, \dots$ ($i = 1, 2$).

By the Weierstrass approximation theorem, we have for any complex-valued continuous function f on $[0, 1]^2$ with period 1 in each variable

$$\int_{[0,1]^2} f(x_1, x_2) dg(x_1, x_2) = \int_{[0,1]^2} f(x_1, x_2) d\tilde{g}_1(x_1, x_2). \quad (4.1)$$

For $0 \leq y < z \leq 1$ and sufficiently small $\delta > 0$, we define the function $f_{yz}^{(\delta)}$ by $f_{yz}^{(\delta)}(x) = 0$ ($0 \leq x \leq y$), $f_{yz}^{(\delta)}(x) = x/\delta - y/\delta$ ($y < x < y + \delta$), $f_{yz}^{(\delta)}(x) = 1$ ($y + \delta \leq x \leq z - \delta$), $f_{yz}^{(\delta)}(x) = -x/\delta + z/\delta$ ($z - \delta < x < z$), $f_{yz}^{(\delta)}(x) = 0$ ($z \leq x \leq 1$).

Since $g(x_1, 1)$ is non-decreasing with respect to x_1 , it has a countable number of dis-continuities $\alpha_{1j} \in [0, 1]$, $j = 1, 2, \dots$. Similarly $\tilde{g}(x_1, 1)$ has a countable number of dis-continuities $\tilde{\alpha}_{1,j} \in [0, 1]$, $j = 1, 2, \dots$. Since $g(1, x_2)$ is non-decreasing with respect to x_2 , it has a countable number of dis-continuities $\alpha_{2j} \in [0, 1]$, $j = 1, 2, \dots$. Similarly $\tilde{g}(1, x_2)$ has a countable number of dis-continuities $\tilde{\alpha}_{2j} \in [0, 1]$, $j = 1, 2, \dots$.

Let $R_1 = (\cup_{j=1}^{\infty} \{(\alpha_{1j}, x_2) : 0 \leq x_2 \leq 1\}) \cup (\cup_{j=1}^{\infty} \{(x_1, \alpha_{2j}) : 0 \leq x_1 \leq 1\})$, $R_2 = (\cup_{j=1}^{\infty} \{(\tilde{\alpha}_{1j}, x_2) : 0 \leq x_2 \leq 1\}) \cup (\cup_{j=1}^{\infty} \{(x_1, \tilde{\alpha}_{1j}) : 0 \leq x_1 \leq 1\})$, and $R = R_1 \cup R_2$.

Let $(y_1, y_2), (z_1, z_2)$ be in $[0, 1]^2 - R$ with $0 < y_i < z_i < 1$ for $i = 1, 2$. Set

$$f_{\delta}(x_1, x_2) = f_{y_1 z_1}^{(\delta)}(x_1) \cdot f_{y_2 z_2}^{(\delta)}(x_2).$$

By (4.1) we obtain

$$\int_{[0,1]^2} f_\delta(x_1, x_2) dg(x_1, x_2) = \int_{[0,1]^2} f_\delta(x_1, x_2) d\tilde{g}(x_1, x_2). \tag{4.2}$$

The left-hand side of (4.2) is written in the form

$$\int_{[0,1]^2} f_\delta(x_1, x_2) dg(x_1, x_2) = \sum_{J_1, J_2} \int_{J_1 \times J_2} f_\delta(x_1, x_2) dg(x_1, x_2)$$

where \sum_{J_1, J_2} stands for the sum over all intervals $J_i \in \{[y_i, y_i + \delta], [y_i + \delta, z_i - \delta], [z_i - \delta, z_i]\}$, $i = 1, 2$.
 If $J_i \neq [y_i + \delta, z_i - \delta]$ for some i , then

$$\lim_{\delta \rightarrow 0} \int_{J_1 \times J_2} f_\delta(x_1, x_2) dg(x_1, x_2) = 0,$$

for in the case of $J_1 = [y_1, y_1 + \delta]$ we have

$$0 \leq \int_{J_1 \times J_2} f_\delta(x_1, x_2) dg(x_1, x_2) \leq \int_{J_1 \times [0,1]} dg(x_1, x_2) = g(y_1 + \delta, 1) - g(y_1, 1) \rightarrow 0$$

as $\delta \rightarrow 0$, since g is continuous at $(y_1, 1)$ by Lemma C.

Furthermore, since g is continuous at (x_1, x_2) with $x_i \in \{y_i, z_i\}$, $i = 1, 2$, we obtain

$$\lim_{\delta \rightarrow 0} \int_{[y_1 + \delta, z_1 - \delta] \times [y_2 + \delta, z_2 - \delta]} f_\delta(x_1, x_2) dg(x_1, x_2) = \Delta_{z_2 - y_2}^{(2)} \Delta_{z_1 - y_1}^{(1)} g(y_1, y_2),$$

hence

$$\lim_{\delta \rightarrow 0} \int_{[0,1]^2} f_\delta(x_1, x_2) dg(x_1, x_2) = \Delta_{z_2 - y_2}^{(2)} \Delta_{z_1 - y_1}^{(1)} g(y_1, y_2).$$

Similarly we obtain

$$\lim_{\delta \rightarrow 0} \int_{[0,1]^2} f_\delta(x_1, x_2) d\tilde{g}(x_1, x_2) = \Delta_{z_2 - y_2}^{(2)} \Delta_{z_1 - y_1}^{(1)} \tilde{g}(y_1, y_2).$$

Therefore from (4.2) it follows that

$$\Delta_{z_2 - y_2}^{(2)} \Delta_{z_1 - y_1}^{(1)} g(y_1, y_2) = \Delta_{z_2 - y_2}^{(2)} \Delta_{z_1 - y_1}^{(1)} \tilde{g}(y_1, y_2). \tag{4.3}$$

Let $(y_{n_1}^{(1)}, y_{n_2}^{(2)})$, $(z_{n_1}^{(1)}, z_{n_2}^{(2)})$ be in $[0, 1]^2 - R$ for $n_i = 1, 2, \dots (i = 1, 2)$ with $0 < y_{i, n_i}^{(i)} < z_{i, n_i}^{(i)} < 1$, $\lim_{n_i \rightarrow \infty} y_{n_i}^{(i)} = 0$, and $\lim_{n_i \rightarrow \infty} z_{n_i}^{(i)} = 1$ for $i = 1, 2$.

From Lemma C, $g(1, 1) = 1$, and $g(\mathbf{x}) = 0$ for any \mathbf{x} with a vanishing coordinate we obtain

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1,2}} \Delta_{z_{n_2}^{(2)} - y_{n_2}^{(2)}}^{(2)} \Delta_{z_{n_1}^{(1)} - y_{n_1}^{(1)}}^{(1)} g(y_{n_1}^{(1)}, y_{n_2}^{(2)}) = 1.$$

From (4.3) we have

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1,2}} \Delta_{z_{n_2}^{(2)} - y_{n_2}^{(2)}}^{(2)} \Delta_{z_{n_1}^{(1)} - y_{n_1}^{(1)}}^{(1)} \tilde{g}(y_{n_1}^{(1)}, y_{n_2}^{(2)}) = 1.$$

Using Lemma 3, we obtain

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1,2}} \tilde{g}(z_{n_1}^{(1)}, z_{n_2}^{(2)}) = 1, \quad \lim_{\substack{n_i \rightarrow \infty \\ i=1,2}} \tilde{g}(t_{n_1}^{(1)}, t_{n_2}^{(2)}) = 0,$$

where $t_{n_j}^{(j)} = y_{n_j}^{(j)}$ for some j and either $t_{n_i}^{(i)} = y_{n_i}^{(i)}$ or $t_{n_i}^{(i)} = z_{n_i}^{(i)}$ for $i = 1, 2$. Since $\tilde{g}(x_1, x_2)$ is non-decreasing with respect to each variable, we have

$$\lim_{\substack{n_i \rightarrow \infty \\ i=1,2}} \tilde{g}(t_{n_1}^{(1)}, t_{n_2}^{(2)}) = 0, \tag{4.4}$$

where $t_{n_j}^{(j)} = y_{n_j}^{(j)}$ for some j and either $t_{n_i}^{(i)} = y_{n_i}^{(i)}$ or $t_{n_i}^{(i)} = z_i$ for $i = 1, 2$.

From (4.3) we obtain

$$\Delta_{z_2 - y_{n_2}^{(2)}}^{(2)} \Delta_{z_1 - y_{n_1}^{(1)}}^{(1)} g(y_{n_1}^{(1)}, y_{n_2}^{(2)}) = \Delta_{z_2 - y_{n_2}^{(2)}}^{(2)} \Delta_{z_1 - y_{n_1}^{(1)}}^{(1)} \tilde{g}(y_{n_1}^{(1)}, y_{n_2}^{(2)}),$$

so

$$\begin{aligned} & \tilde{g}(z_1, z_2) - \tilde{g}(z_1, y_{n_2}^{(2)}) - \tilde{g}(y_{n_1}^{(1)}, z_2) + \tilde{g}(y_{n_1}^{(1)}, y_{n_2}^{(2)}) \\ &= g(z_1, z_2) - g(z_1, y_{n_2}^{(2)}) - g(y_{n_1}^{(1)}, z_2) + g(y_{n_1}^{(1)}, y_{n_2}^{(2)}). \end{aligned}$$

By using (4.4) and continuity of g at $(z_1, 0)$, $(0, z_2)$. we get

$$\tilde{g}(z_1, z_2) = g(z_1, z_2).$$

Letting $z_1 \rightarrow 1$, we get

$$\lim_{z_1 \rightarrow 1} \tilde{g}(z_1, z_2) = g(1, z_2), \tag{4.5}$$

because g is continuous at $(1, z_2)$.

Set $z_1 = z_{n_1}^{(1)}$, $y_2 = y_{n_2}^{(2)}$ in (4.3), then by letting $n_1 \rightarrow \infty, n_2 \rightarrow \infty$ we have

$$\lim_{n_1 \rightarrow \infty} \tilde{g}(z_{n_1}^{(1)}, z_2) - \tilde{g}(y_1, z_2) = g(1, z_2) - g(y_1, z_2).$$

Using (4.5), we have

$$\tilde{g}(y_1, z_2) = g(y_1, z_2).$$

In the same way, we can show that $\tilde{g}(x_1, x_2) = g(x_1, x_2)$ for $x_i \in \{y_i, z_i\}$ ($i = 1, 2$). Therefore we obtain

$$G(\mathbf{a}_n \bmod 1) \subset G(\tilde{\mathbf{a}}_n \bmod 1). \tag{4.6}$$

If we assume that $\tilde{g}(x_1, 1)$ and $\tilde{g}(1, x_2)$ is continuous at $x_i = 0, 1$ for $i = 1, 2$, then (4.6) is shown in the same way, hence

$$G(\tilde{\mathbf{a}}_n \bmod 1) \subset G(\mathbf{a}_n \bmod 1).$$

Thus we have

$$G(\mathbf{a}_n \bmod 1) = G(\tilde{\mathbf{a}}_n \bmod 1).$$

This proof can be extended to the general s -dimensional case.

Acknowledgements. The author thanks an anonymous referee for a careful reading of this paper and several useful remarks.

REFERENCES

- [1] AKIYAMA, S.: *Almost uniform distribution modulo 1 and the distribution of primes*, Acta Math. Hungar. **78** (1998), no.1-2, 39-44.
- [2] BRUNK, H. D. – EWING, G. M. – UTZ, W. R.: *Some Helly theorems for monotone functions*, Proc. Amer. Math. Soc. **7** (1956), no.5, 776-783.
- [3] GIULIANO ANTONINI, R. – STRAUCH, O.: *On weighted distribution function of sequences*, Uniform Distribution Theory **3** (2008), no. 1, 1-18.
- [4] GRABNER, P.S. – STRAUCH, O. – TICHY, R. F.: *L^p -discrepancy and statistical independence of sequences*, Czechoslovak Math. J. **49(124)** (1999), no. 1, 97-110.
- [5] GRABNER, P. J. – TICHY, R. F.: *Remarks on statistical independence of sequences*, Math. Slovaca **44** (1994), no.1, 91-94.
- [6] KUIPERS, L. – NIEDERREITER, H.: *Uniform distribution of Sequences*, Pure and Applied Mathematics, John Wiley & Sons, New York, London, Sydney, Tronto, 1974.
- [7] KOKSMA, J. F.: *Asymptotische verdeling van gattallen modulo 1. I,II,III*, Mathematica (Leiden) **1** (1933), 1-6, 107-114.
- [8] RAUZY, G.: *Propriétés statistiques de suites arithmétiques*, Le Mathématicien, Vol.15, Collection SUP, Presses Universitaires de France, Paris, 1976, 133pp.
- [9] RIBENBOIM, P.: *The New Book of Prime Number Records*, Springer-Verlag, New York, 1995.
- [10] STRAUCH, O.: *L^2 discrepancy*, Math. Slovaca **44** (1994), no.5,601-632.
- [11] STRAUCH, O.: *Uniformly maldistributed sequences in a strict sense*, Mh. Math. **120** (1995), 153-164.
- [12] STRAUCH, O. – BLAŽEKOVÁ, O.: *Distribution of the sequence $p_n/n \bmod 1$* , Uniform Distribution Theory **1** (2006), no. 1, 45-63.
- [13] STRAUCH, O. – NAIR, R.: *Unsolved Problems* (version date June 29, 2011), <http://www.boku.ac.at/MATH/udt/unsolvedproblems.pdf>.
- [14] STRAUCH, O. – PORUBSKÝ, S.: *Distribution of Sequences: A Sampler*, Schriftenreihe der Slowakischen Akademie der Wissenschaften, Band 1, Peter Lang, Bern, Frankfurt and Main, 2005.
- [15] VINOGRADOV, I. M.: *The Method of Trigonometric Sums in the Theory of Numbers*, Interscience, London-New York, 1954.

YUKIO OHKUBO

- [16] WINTNER, A.: *On the cyclical distribution of the logarithms of the prime numbers*, Quart. J. Math. (1) **6** (1935), 65–68.

Received January 24, 2011
Accepted November 15, 2011

Yukio Ohkubo
Department of Business Administration
The International University of Kagoshima
8-34-1 Sakanoue, Kagoshima-shi, 891-0197
JAPAN
E-mail: ohkubo@eco.iuk.ac.jp