

## Computing a Nearest Correlation Matrix with Factor Structure

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Joint work with Rüdiger Borsdorf, Marcos Raydan

Linear Algebra and its Applications to Financial Engineering—in honour of Prof. Peter Lancaster.

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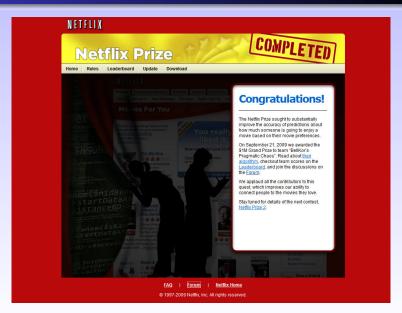
## http://www.movielens.org



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#### **Correlation Matrix**

 $n \times n$  symmetric positive semidefinite matrix A with  $a_{ii} \equiv 1$ .

- symmetric,
- 1s on the diagonal,
- eigenvalues nonnegative or all principal minors nonnegative.

#### **Properties:**

- off-diagonal elements between −1 and 1,
- convex set.

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## Quiz

#### Is this a correlation matrix?

$$\begin{bmatrix} 1 & 1 & 0 \\ 1 & 1 & 1 \\ 0 & 1 & 1 \end{bmatrix}.$$

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#### Is this a correlation matrix?

$$\begin{bmatrix} 1 & 1 & 0 \\ 1 & 1 & 1 \\ 0 & 1 & 1 \end{bmatrix}. \qquad \text{Spectrum: } -0.4142, \, 1.0000, \, 2.4142.$$

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#### Is this a correlation matrix?

#### For what w is this a correlation matrix?

$$\begin{bmatrix} 1 & w & w \\ w & 1 & w \\ w & w & 1 \end{bmatrix}.$$

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#### Is this a correlation matrix?

#### For what w is this a correlation matrix?

$$\begin{bmatrix} 1 & w & w \\ w & 1 & w \\ w & w & 1 \end{bmatrix} \cdot \frac{-1}{n-1} \le w \le 1.$$

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### **Structured Correlation Matrices**

Nonnegative:

$$\begin{bmatrix} 1 & \frac{1}{2} & \frac{1}{3} \\ \frac{1}{2} & 1 & \frac{1}{4} \\ \frac{1}{3} & \frac{1}{4} & 1 \end{bmatrix}.$$

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#### **Structured Correlation Matrices**

■ Nonnegative:

$$\begin{bmatrix} 1 & \frac{1}{2} & \frac{1}{3} \\ \frac{1}{2} & 1 & \frac{1}{4} \\ \frac{1}{3} & \frac{1}{4} & 1 \end{bmatrix}.$$

Low rank:

$$\begin{bmatrix} 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{bmatrix}.$$

#### **Structured Correlation Matrices**

■ Nonnegative:

$$\begin{bmatrix} 1 & \frac{1}{2} & \frac{1}{3} \\ \frac{1}{2} & 1 & \frac{1}{4} \\ \frac{1}{3} & \frac{1}{4} & 1 \end{bmatrix}.$$

Low rank:

$$\begin{bmatrix} 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{bmatrix}.$$

Factor structure:

$$\begin{bmatrix} 1 & x_1x_2 & x_1x_3 \\ x_1x_2 & 1 & x_2x_3 \\ x_1x_3 & x_2x_3 & 1 \end{bmatrix}.$$

## **Approximate Correlation Matrices**

Empirical correlation matrices often not true correlation matrices, due to

- asynchronous data
- missing data
- limited precision
- stress testing



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#### **Nearest Correlation Matrix**

#### Find X achieving

$$\min\{\|A - X\|_F : X \text{ is a correlation matrix }\},$$

where 
$$||A||_F^2 = \sum_{i,j} a_{ij}^2$$
.

- ★ Constraint set is a closed, convex set, so unique minimizer.
- ★ Nonlinear optimization problem.

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#### **Questions From Finance Practitioners**

"Given a real symmetric matrix A which is almost a correlation matrix what is the best approximating (in Frobenius norm?) correlation matrix?"

"I am researching ways to make our company's correlation matrix positive semi-definite."

"Currently, I am trying to implement some real options multivariate models in a simulation framework. Therefore, I estimate correlation matrices from inconsistent data set which eventually are non psd."

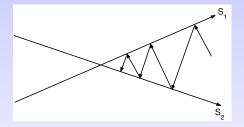
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Result Page:

## **Alternating Projections Method**

H (2002): repeatedly **project** onto the positive semidefinite matrices then the unit diagonal matrices.



- Easy to implement.
- ► Guaranteed convergence, at a linear rate.
- ► Can add further constraints/projections, e.g., fixed elements (Lucas, 2001).

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#### **Newton Method**

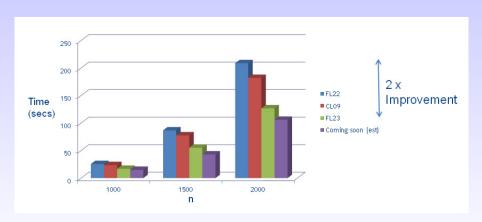
Qi & Sun (2006): **Newton method** based on theory of strongly semismooth matrix functions.

- Applies Newton to **dual** (unconstrained) of min  $\frac{1}{2}||A X||_F^2$  problem.
- Globally and quadratically convergent.
- H & Borsdorf (2010) improve efficiency and reliability:
  - use minres with preconditioning,
  - reliability improved by line search tweaks,
  - extra scaling step to ensure unit diagonal.

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## **NAG** Library

Implemented in NAG codes g02aaf (g02aac) and g02abf (weights, lower bound on ei'vals—Mark 23).



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#### **Factor Model**

$$\xi = \underbrace{X}_{n \times k} \underbrace{\eta}_{k \times 1} + \underbrace{F}_{n \times n} \underbrace{\varepsilon}_{n \times 1}, \qquad \eta_i, \varepsilon_i \in N(0, 1),$$

where  $F = diag(f_{ii})$ . Since  $E(\xi) = 0$ ,

$$cov(\xi) = E(\xi \xi^T) = XX^T + F^2.$$

Assume  $\operatorname{var}(\xi_i) \equiv 1$ . Then  $\sum_{j=1}^k x_{ij}^2 + f_{ii}^2 = 1$ , so

$$\sum_{i=1}^{k} x_{ij}^2 \le 1, \qquad i = 1: n.$$

- Collateralized debt obligations (CDOs),
- multifactor normal copula model.

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#### Structured Correlation Matrix

Yields correlation matrix of form

$$C(X) = D + XX^{T} = D + \sum_{j=1}^{k} x_{j}x_{j}^{T},$$

$$D = \operatorname{diag}(I - XX^{T}), \qquad X = [x_{1}, \dots, x_{k}].$$

C(X) has k factor correlation matrix structure.

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#### Structured Correlation Matrix

Yields correlation matrix of form

$$C(X) = D + XX^{T} = D + \sum_{j=1}^{k} x_{j}x_{j}^{T},$$
 $D = \text{diag}(I - XX^{T}), \qquad X = [x_{1}, \dots, x_{k}].$ 

C(X) has k factor correlation matrix structure.

$$C(X) = \begin{bmatrix} 1 & y_1^T y_2 & \dots & y_1^T y_n \\ y_1^T y_2 & 1 & \dots & \vdots \\ \vdots & & \ddots & y_{n-1}^T y_n \\ y_1^T y_n & \dots & y_{n-1}^T y_n & 1 \end{bmatrix}, \quad y_i \in \mathbb{R}^k.$$

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#### Aims

For *k* factor correlation matrices, investigate

- mathematical properties,
- nearness problem.

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#### 1-Parameter Correlation Matrix

$$X(w) = \begin{bmatrix} 1 & w & w \\ w & 1 & w \\ w & w & 1 \end{bmatrix}, \qquad w \in \mathbb{R}.$$

#### Theorem

 $\min\{\|A - X(w)\|_F : X(w) \text{ a corr. matrix}\}\$ has unique solution the projection of

$$w = \frac{e^T A e - \operatorname{trace}(A)}{n^2 - n},$$

onto [-1/(n-1), 1].

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#### **Block Structured Correlation Matrix**

$$\begin{bmatrix} 1 & \gamma_{11} & \gamma_{12} & \gamma_{12} \\ \gamma_{11} & 1 & \gamma_{12} & \gamma_{12} \\ \hline \gamma_{12} & \gamma_{12} & 1 & \gamma_{22} \\ \gamma_{12} & \gamma_{12} & \gamma_{22} & 1 \end{bmatrix}, \quad \boldsymbol{C_{ij}} = \begin{cases} \boldsymbol{C}(\gamma_{ii}) \in \mathbb{R}^{n_i \times n_i}, & i = j, \\ \gamma_{ij} \boldsymbol{e} \boldsymbol{e}^T \in \mathbb{R}^{n_i \times n_j}, & i \neq j. \end{cases}$$

#### Objective function:

$$f(\Gamma) = \|A - C(\Gamma)\|_F^2 = \sum_{i=1}^m \|A_{ii} - C(\gamma_{ii})\|_F^2 + \sum_{i \neq j} \|A_{ij} - \gamma_{ij}ee^T\|_F^2.$$

- Convex constraint set ⇒ unique minimizer.
- Alternating projections converges (use prev. theorem for projection onto pattern).

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#### 1-Factor Correlation Matrix

$$C(x) = diag(1 - x_i^2) + xx^T, \qquad x \in \mathbb{R}^n$$

i.e.,  $c_{ij} = x_i x_j$ ,  $i \neq j$ .

#### Lemma

$$\det(C(x)) = \prod_{i=1}^{n} (1 - x_i^2) + \sum_{i=1}^{n} x_i^2 \prod_{\substack{j=1 \ j \neq i}}^{n} (1 - x_j^2).$$

#### Corollary

If  $|x| \le e$  with  $x_i = 1$  for at most one i then C(x) is nonsingular. C(x) is singular if  $x_i = x_i = 1$  for some  $i \ne j$ .

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#### Rank Result

$$C(x) = \operatorname{diag}(1 - x_i^2) + xx^T,$$

#### Theorem

Let  $x \in \mathbb{R}^n$  with  $|x| \le e$ . Then  $\operatorname{rank}(C(x)) = \min(p+1, n)$ , where p is the number of  $x_i$  for which  $|x_i| < 1$ .

$$x = \begin{bmatrix} 1 & 1 & 1 & x_4 & x_5 \\ 1 & 1 & 1 & x_4 & x_5 \\ 1 & 1 & 1 & x_4 & x_5 \\ 1 & 1 & 1 & x_4 & x_5 \\ x_4 & x_4 & x_4 & 1 & x_4 x_5 \\ x_5 & x_5 & x_5 & x_4 x_5 & 1 \end{bmatrix}.$$

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## All your hedges in one basket

Leif Andersen, Jakob Sidenius and Susanta Basu present new techniques for single-tranche CDO sensitivity and hedge ratio calculations. Using factorisation of the copula correlation matrix, discretisation of the conditional loss distribution followed by a recursion-based probability calculation, and derivation of analytical formulas for deltas, they demonstrate a significant improvement in computational speeds

n a traditional synthetic collateralised debt obligation (CDO), the arranger tranches out credit losses on a pool of credit default swaps (CDSs) and passes them through to different investors. Assuming that investors for all tranches can be identified, the arranger is typically left with fairly moderate market exposure. For various reasons, placing the entire pool capital structure with investors has become increasingly difficult, and many recent credit basket derivatives expose the dealer to significant market risk. For instance, the recent 'single-tranche' CDO (STCDO) product involves the sale of a single CDO tranche to a single customer, leaving it to the arranger to manage the risk of the remaining capital structure. As STCDOs and similar 'custom' products offer significant customer benefits and are much less difficult to originate than traditional CDOs, such products are likely to increase in importance. This is especially true for managed trades where the customer has certain rights to alter the composition of the reference portfolio over time.

A basic prerequisite for active management of the risk of a credit basket derivative is the ability to accurately calculate the sensitivity of the security with respect to market and model parameters, most prominently the par CDS spreads of the underlying reference pool. The numbers of such sensitivities can be very large - many thousands - and can put considerable strain on computing resources. Moreover, the calculation of each of where Q is the risk-neutral probability measure and  $\lambda_k$  is a (forward) default hazard rate function. The functions  $p_k(T)$ , k = 1, ..., N can be bootstrapped by standard means from the quoted CDS spreads and are assumed known for all T.

Equation (1) fully establishes the risk-neutral marginal distribution of each default time t.. To construct the joint distribution of all default times, we here choose2 to employ a Student-t copula, which we quickly define for reference. Defining vectors  $\mathbf{\tau} = (\tau_1, \dots, \tau_N)^T$  and  $\mathbf{T} = (T_1, \dots, T_N)^T$ , the joint default time distribution in the Student-t copula, becomes:

$$Q(\tau \leq T) = t_{N,v}(t_{1,v}^{-1}(p_1(T_1)),...,t_{1,v}^{-1}(p_N(T_N)))$$
 (2)

where  $t_1$ , and  $t_N$ , are the one- and N-dimensional cumulative Student-t distribution functions with v degrees of freedom, respectively. Recall that the density  $\eta_N$  of an N-dimensional Student-t distribution with correlation matrix \( \Sigma \) is:

$$\eta_{N,v}(\mathbf{z}) = C_{N,v} \left( 1 + v^{-1} \mathbf{z}^T \Sigma^{-1} \mathbf{z} \right)^{-\frac{v_1 N}{2}}, \quad C_{N,v} = \frac{\Gamma\left(\frac{v_1 + N}{2}\right)}{\Gamma\left(\frac{v_2}{2}\right) \sqrt{\left|\Sigma\right| \left(v \pi\right)^N}}$$
(3)

where  $\Gamma$  is the gamma function. For high degrees of freedom, (3) approaches

#### **One-Factor Problem**

$$\min_{x \in \mathbb{R}^n} f(x) := \|A - C(x)\|_F^2$$
  
subject to  $-e \le x \le e$ .

- Objective function is nonconvex.
- The constraint implies C(x) is a correlation matrix.

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#### One-Factor Problem: Derivatives

Objective:

$$f(x) = \langle A - I, A - I \rangle_F - 2x^T (A - I)x + (x^T x)^2 - \sum_{i=1}^n x_i^4$$

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Gradient:

$$\nabla f(x) = 4((x^T x)x - (A - I)x - \operatorname{diag}(x_i^2)x).$$

• Hessian:

$$\nabla^2 f(x) = 4(2xx^T + (x^Tx + 1)I - A - 3\text{diag}(x_i^2)).$$

- $\nabla f(x)$ ,  $\nabla^2 f(x)$  cheap.
- f(x) has a saddle point at x = 0.

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#### k Factor Problem

$$C(X) := I - \operatorname{diag}(XX^T) + XX^T$$
 with  $X \in \mathbb{R}^{n \times k}$ .

Representation not unique!

$$\sum_{i=1}^k x_{ij}^2 \le 1 \quad \Longrightarrow \quad C(X) \text{ is a correlation matrix.}$$

The k factor problem is

$$\min_{X \in \mathbb{R}^{n \times k}} f(x) := \|A - C(X)\|_F^2$$
 subject to  $\sum_{j=1}^k x_{ij}^2 \le 1$ .

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#### k Factor Problem: Derivatives

Gradient

$$\nabla f(X) = 4(X(X^TX) - AX + X - \operatorname{diag}(XX^T)X)$$

• Hessian given implicitly, can be viewed as a matrix representation of the Fréchet derivative of  $\nabla f(X)$ .

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## **Choice of Optimization Method**

- Derivatives available.
- Ignore the constraints?
- Starting matrix, convergence test?

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## **Choice of Optimization Method**

- Derivatives available.
- Ignore the constraints?
- Starting matrix, convergence test?
- Rich set of solvers in NAG Library, Mark 22:
  - E04 Minimizing or Maximizing a Function
  - E05 Global Optimization of a Function
- MATLAB Optimization toolbox.

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## **Alternating Directions**

$$f(x_{ij}) = \text{const.} + 2\sum_{q \neq i} \left(a_{iq} - \sum_{s=1}^k x_{is}x_{qs}\right)^2.$$

Hence  $f'(x_{ij}) = 0$  if

$$x_{ij} = \frac{\sum_{q \neq i} x_{qj} \left( a_{iq} - \sum_{s \neq j} x_{is} x_{qs} \right)}{\sum_{q \neq i} x_{qj}^2}.$$

Project  $x_{ij}$  onto [-1, 1].

- Convergence to stationary point of *f* guaranteed.
- Limit may not be feasible for k > 1.

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## **Principal Factors Method**

Anderson, Sidenius & Basu (2003): with 
$$F(X) = I - \text{diag}(XX^T)$$
 [so  $C(X) = F(X) + XX^T$ ]

$$X_i = \operatorname{argmin}_{X \in \mathbb{R}^{n \times k}} \|A - F(X_{i-1}) - XX^T\|_F.$$

Min obtained by eigendecomposition of  $A - F(X_{i-1})$ . Equivalent to alternating projections method for

$$\mathcal{U} := \{ W \in \mathbb{R}^{n \times n} : w_{ij} = a_{ij} \text{ for } i \neq j \}$$
 convex,  
 $\mathcal{S} := \{ W \in \mathbb{R}^{n \times n} : W = XX^T \text{ for } X \in \mathbb{R}^{n \times k} \}$  nonconvex!

- Alt proj theory says no guarantee of convergence!
- Constraints ignored, so project final iterate onto them.

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## Spectral Projected Gradient Method

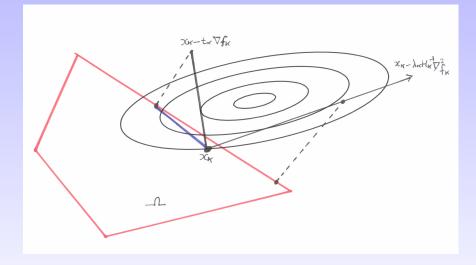
Birgin, Martínez & Raydan (2000).

To minimize  $f: \mathbb{R}^n \to \mathbb{R}$  over convex set  $\Omega$ :

$$\mathbf{X}_{k+1} = \mathbf{X}_k + \alpha_k \mathbf{d}_k.$$

- $d_k = P_{\Omega}(x_k t_k \nabla f(x_k)) x_k$  is descent direction,
- $\alpha_k \in [-1, 1]$  chosen through nonmonotone line search strategy.
- Promising since  $P_{\Omega}(\cdot)$  is cheap to compute.

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- Fortran Code in ACM TOMS (Alg 813).
- R package by Varadhan & Gilbert (2009).

## Test Examples

- Orr: gallery('randcorr',n)
- **nrand**:  $\frac{1}{2}(B+B^T) + \text{diag}(I-B)$  with  $B \in [-1,1]^{n \times n}$  such that  $\lambda_{\min}(B) < 0$ .

Results averaged over 10 instances.

- AD: alternating directions.
- PFM: principal factors method.
- Nwt: e041b of NAG Toolbox for MATLAB (modified Newton), bound constraints.
- **SPGM**: spectral projected gradient method.

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## Comparison: k = 1, n = 2000

	tol= 10 <sup>-3</sup>			tol= 10 <sup>-6</sup>					
	t(sec.)	#its	$\sqrt{f(X^*)}$	t(sec.)	#its	$\sqrt{f(X^*)}$			
<b>corr</b> , $f(X_0) = 26.0$									
AD	3.3	5.2	26.0	3938	7282	26.0			
PFM	68	1.1	26.0	827	18	26.0			
Nwt	23	1.8	26.0	36	5.0	26.0			
SPGM	9.8	5.2	26.0	638	760	26.0			
<b>nrand</b> $f(X_0) = 825.13$									
AD	3.8	7.2	815.79	3.4	10.0	815.79			
PFM	22	3.0	815.81	19.0	4.0	815.81			
Nwt	4167	1222	815.79	4312	1229	815.79			
SPGM	9.4	7.2	815.79	11	9.6	815.79			

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## Comparison: k = 6, n = 1000

	t	<b>ol</b> = 10 <sup>-</sup>	-3	<b>tol</b> = 10 <sup>-6</sup>					
	t(sec.)	#its	$\sqrt{f(X^*)}$	t(sec.)	#its	$\sqrt{f(X^*)}$			
<b>corr</b> , $f(X_0) = 18.5$									
AD	704	836	18.38	5060	5955	18.38			
PFM	10	4.1	18.38	95	28.1	18.38			
Nwt	167	52	18.38	280	68.2	18.38			
SPGM	24	235	18.38	108	892	18.38			
<b>nrand</b> , $f(X_0) = 415$									
AD	8694	9816	421	1.13e4	1.28e4	414			
PFM	10.1	6.0	421	9.8	10	420			
Nwt	146	40.8	421	109	56	420			
SPGM	122	1263	407	276	2925	407			

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#### **Overall Conclusions**

- Nearest k factor correlation matrix problem relevant in many applications.
- Principal factors method has no convergence theory and can converge to an incorrect answer!
- Important to use methods that respect the constraints and converge to a feasible stationary point.
- Spectral projected gradient method is best choice: exploits convexity of constraints. Implemented in NAG routine g02aef (Mark 23, 2011).

R. Borsdorf, N. J. Higham and M. Raydan Computing a Nearest Correlation Matrix with Factor Structure, SIMAX, 31(5): 2603–2622, 2010.

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